

Sparse Prototype Code Underlies Classification and Prediction Across Modalities

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Abstract

Neural representations have become a central tool for studying the internal mechanisms of modern AI models, yet their complex high-dimensional structure makes them difficult to interpret. We show that classification tasks give rise to a universal representational geometry, shared across state-of-the-art models in vision, audio, and language processing. The key structure is that within-class variability is not random in representation space. Instead, its classifier-relevant component has strong and structured correlations with the class's own centroid and with the centroids of its competing classes. Building on this observation, we derive an analytical mean-field theory governed mainly by the variability along true-class and rival-class centroid coordinates, together with a global renormalization of the class radius that compensates for the non-Gaussian statistics of real representations. The theory accurately predicts classification accuracy across architectures and modalities. The relevant geometric quantities improve systematically with model scale, mirroring the observed gains in accuracy. A striking feature of the theory is its sparsity: accurate prediction requires only a small set of centroid coordinates associated with the true class and its strongest rivals - connecting our framework to sparse-feature extraction approaches such as sparse autoencoders. Together, these results provide a parsimonious predictive theory of neural representations and suggest that classification in deep networks is governed by a sparse, centroid-aligned structure embedded within the full high-dimensional representation space.

1 Introduction

One of the central challenges in modern Deep Neural Network (DNN) research is to open the "black box" of computation and to understand the internal mechanisms by which networks transform their inputs into useful predictions. Such understanding is essential for building safer, more reliable AI systems [Bommasani et al., 2021, Hendrycks et al., 2021]. A primary tool toward this goal is the analysis of representations - the hidden-layer activations induced by specific inputs. These representations form structured point clouds whose geometry reflects both the data and the model, but extracting interpretable principles from this structure remains a significant challenge.

A growing body of work in mechanistic interpretability addresses this challenge by extracting sparse, human-interpretable features from network activations [Elhage et al., 2022, Bricken et al., 2023, Huben et al., 2024, Templeton et al., 2026]. Sparse autoencoders, feature dictionaries, and circuit analyses have shown that semantically meaningful concepts are often encoded along specific directions in activation space and can be isolated within otherwise opaque representations. These approaches are powerful tools for identifying the essential features of the network; they are typically not aimed, however, at providing a quantitative theory linking the geometric structure of a representation to how well it supports a downstream task. Such a theory would help address fundamental questions, such as why some networks generalize better than others.

Predictive theories of representational geometry —manifold capacity [Chung et al., 2018, Cohen et al., 2020, Stephenson et al., 2021, Wakhloo et al., 2023] and neural collapse [Papayan et al., 2020, Han et al., 2021, Galanti et al., 2021] — have made important progress in connecting geometry to performance within their respective regimes. Capacity theory studies the linear separability of object manifolds under random binary dichotomies, relating global capacity to geometric quantities such as manifold radius, effective dimension, and inter-manifold correlations. Neural collapse describes the final phase of supervised training, in which within-class variability collapses and class means converge to a simplex equiangular tight frame. Modern pretrained foundation models, however, typically operate in a different regime: they are evaluated using multiclass accuracy and are often trained with early stopping or self-supervised objectives, leaving substantial within-class variability. Recent work connects pairwise class geometry to performance and generalization bounds in few-shot learning [Sorscher et al., 2022, Luthra et al., 2026]. A predictive geometric theory of multiclass accuracy in the finite-radius regime that accounts for both competition among multiple classes and the rich within-class statistics of many examples is currently missing.

Our analytical lens is prototype-based classification. A long line of work - from nearest-class-mean classifiers [Mensink et al., 2013] to prototypical networks for few-shot learning [Snell et al., 2017] - has established that classifying inputs by proximity to class centroids is a strong strategy whenever representations are well-structured. Adopting the centroid-based classifier in this work has a clear advantage: the centroids are geometric objects, allowing classification accuracy to be related directly to geometric measures of the representation - while remaining relevant for realistic classification settings. Thus, we leverage this framework to close the theoretical gap, deriving and empirically validating a quantitative geometric theory of multi-class classification accuracy using prototype classification, and apply it across architectures and modalities.

Our main contributions are:

Universal centroid-aligned structure: We identify a universal representational geometry shared across vision, language, and audio backbones. Its defining property is that within-class variability is strongly correlated with both the true-class centroid and the centroids of competing classes (see Fig. 1a). We explicitly demonstrate this by surgically removing only these centroid-aligned correlations (using rotations, see Fig. 1b) - leaving centroid-centroid correlations, total variance, and example-example structure intact - and find that doing so drives per-class accuracy to essentially 100%. The centroid-aligned correlations are therefore precisely the geometric source of classification errors.

Sparsity of competition: These correlations do not spread evenly across the other classes; for each class, the geometry concentrates on a small set of active rivals, which is sufficient to recover the full geometric picture and reproduce the predicted per-class accuracy. These rivals also appear to carry semantically meaningful relations to the true class (see Fig. 1a). Each example’s classification is therefore faithfully captured by a sparse, semantically meaningful centroid-aligned basis embedded in the full high-dimensional representation - a grounded analog of the sparse-feature picture pursued in mechanistic interpretability.

Compact predictive mean-field theory: To connect representational geometry to classification, we adopt a prototype classifier - scoring examples by dot product with unit-normalized class centroids [Snell et al., 2017] - which serves as an empirically reliable proxy for the trained linear head (see Appendix A.1). Building on this, we derive an analytical mean-field theory in the sparse centroid basis (Sec. 5). Its inputs are the variability along the true- and rival-class centroid directions and a single global radius renormalization that absorbs the heavy-tailed statistics of real representations (see Fig. 3c). Crucially, this factor is tightly correlated with a direct measure of tail heaviness (Fig. 4e). From these inputs alone, the theory predicts per-class accuracy across classes, architectures, and modalities (see Fig. 4(a-d), Appendix H).

Geometry across model scales: Tracking the theory’s geometric ingredients across model families, we find that they systematically improve as the models scale up, mirroring the gains in accuracy (Table 4). Importantly, we show that larger models achieve their accuracy advantage without having the smallest overall within-class variance. Our theory identifies the specific geometric measures that drive this improvement - taking a step toward understanding the key geometrical ingredients required for good generalization in real state-of-the-art models. We hypothesize that these measures are practically actionable: targeting them during training could improve generalization; matching

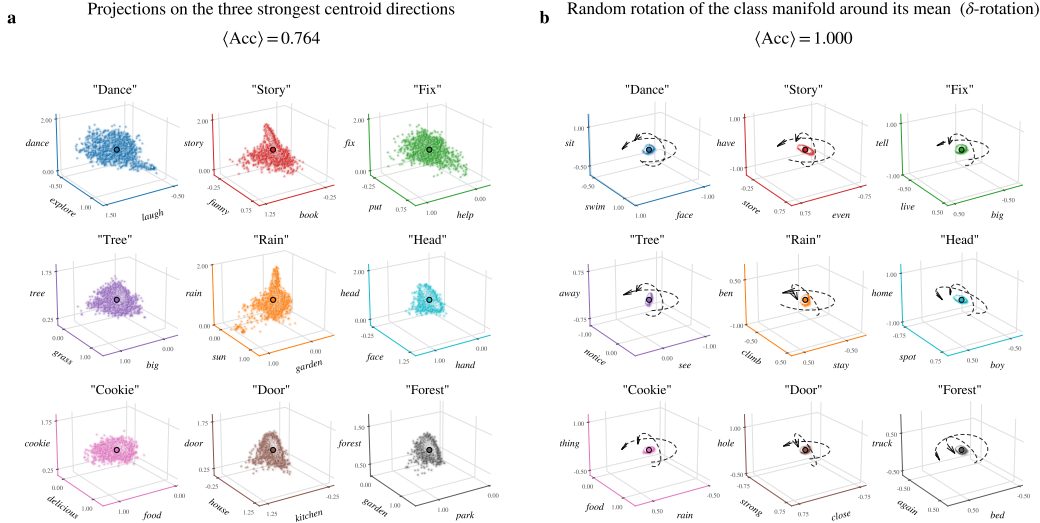


Figure 1: Centroid-aligned variability, Qwen3-32B, TinyStories representations. The variability of each class is highly correlated with other class centroids. **(a)** To visualize it, we use Qwen3-32B representations [Yang et al., 2025] on the dataset TinyStories [Eldan and Li, 2023]. We project each example onto the three centroids with the largest correlations among the class variability. For 217/320 classes (68%), the true class centroid is among these top 3 (the nine shown here are of this sort). The rivals carry semantic meaning (e.g., the rivals of "cookie" are "food" and "delicious"; the rivals of "head" are "face" and "hand"). The mean prototype classification accuracy of these classes (see Def. 2.1) is 76.4%. **(b)** To causally test the role of the variability-centroid correlations, we rotate the class manifold around its mean (denoted as δ -rotation), breaking these correlations while leaving the class radius and the distance between class centroids unchanged. The manipulation removes all classification errors, and the projections onto the top-3 directions shrink to 11.3% of their original magnitude on average, and those directions no longer carry semantic meaning.

them between teacher and student could guide more geometry-aware distillation [Hinton et al., 2015] and could help identify the best-suited pretrained backbone for transfer [Kornblith et al., 2018].

2 Setup and Notation

We study how the geometry of neural representations determines classification accuracy. For each input we consider a representation vector taken from some layer of a neural network, and the task is to assign it to one of P classes.

Prototype classification: Rather than training a linear readout on top of the representation, we use a prototype classifier [Snell et al., 2017], in which the class-specific readout vector is not a learned weight but the centroid - the empirical mean of the class in representation space (Def. 2.1), normalized to a unit vector. Centroids are estimated from a labeled set, and at inference, the learner classifies one example against all P centroids (computed without it, "Leave one out" [Lachenbruch and Mickey, 1968]). This has a direct geometric definition while remaining relevant for realistic classification, can be applied at every layer of every model without additional fine-tuning, and empirically tracks trained linear readouts closely (see Appendix A.1), making it both a flexible evaluator in its own right and a useful proxy for the standard linear head.

Experimental setup: We evaluate frozen representations from pretrained backbones spanning vision (ImageNet-1K [Deng et al., 2009]), language (masked and next-token prediction on TinyStories [Eldan and Li, 2023]), and audio (VGGSound [Chen et al., 2020]), across transformer and convolutional architectures, supervised, self-supervised, and contrastive objectives. Full lists of datasets, models, and extraction details are in Appendix B. For language, the task is to predict a content word from its representation. In autoregressive models this is the embedding of the token preceding the word; in bidirectional models it is the embedding at the word's masked position. We group each word's inflections under a single lemma label (e.g. say/says/said, friend/friends), and keep only lemmas with at least 1000 examples, so their statistics are well estimated. Full technical details are in Appendix B.

Definition 2.1 (*prototype classification*): Let $\mathbf{x} \in \mathbb{R}^N$ be the representation of an input at a given layer of a neural network, and consider a P -class task with labels $y \in \{1, \dots, P\}$ and dataset $\mathcal{D} = \{(\mathbf{x}_i, y_i)\}_{i=1}^n$. For each class μ , define the centroid $\mathbf{c}_\mu = \mathbb{E}[\mathbf{x} \mid y = \mu]$ and its unit direction $\hat{\mathbf{c}}_\mu = \mathbf{c}_\mu / \|\mathbf{c}_\mu\|$, which serves as the prototype class readout. Throughout the paper, we use the terms centroid and prototype interchangeably. The prototype logits and accuracy are defined as

$$t_\mu(\mathbf{x}) = \hat{\mathbf{c}}_\mu \cdot \mathbf{x}, \quad \text{Acc} = \mathbb{E}_{(\mathbf{x}, y) \in \mathcal{D}} [\mathbf{1}\{\arg \max_\mu t_\mu = y\}]. \quad (1)$$

Definition 2.2 (*manifold geometry*): The class manifold is defined as the point cloud of a particular class k : $\mathcal{M}_k = \{\mathbf{x}_i : y_i = k\}$. Each $\mathbf{x} \in \mathcal{M}_k$ decomposes as $\mathbf{x} = \mathbf{c}_k + \delta\mathbf{x}$ with $\mathbb{E}_{\mathbf{x} \in \mathcal{M}_k}[\delta\mathbf{x}] = \mathbf{0}$. We name $\delta\mathbf{x}$ the **residual** vector. The class radius R measures the overall within-class variance, and $\sigma_\mu^2 \in [0, 1]$ is the fraction of the residual variance along a particular centroid direction $\hat{\mathbf{c}}_\mu$:

$$R^2 = \frac{\mathbb{E}_k[\|\delta\mathbf{x}\|^2]}{\|\mathbf{c}_k\|^2}, \quad \sigma_\mu^2 = \frac{\mathbb{E}_k[(\delta\mathbf{x} \cdot \hat{\mathbf{c}}_\mu)^2]}{\mathbb{E}_k[\|\delta\mathbf{x}\|^2]}. \quad (2)$$

Throughout, geometric quantities depend on a fixed true class k . We treat k as fixed and drop the subscript to lighten notation: $R \equiv R_k$, $\sigma_{k\mu} \equiv \sigma_\mu$, and σ_k^2 is the fraction of residual variance along the true centroid $\hat{\mathbf{c}}_k$. In addition, $\mathbb{E}_k$ denotes the average over all the points in the manifold k , $\mathbb{E}_k \equiv \mathbb{E}_{\mathbf{x} \in \mathcal{M}_k}$.

3 Isotropic Uncorrelated Model

We first examine a baseline of an uncorrelated model, in which the centroids $\mathbf{c}_\mu$ and the residuals $\delta\mathbf{x}$ are drawn independently from an isotropic Gaussian distribution. The model yields testable predictions for how geometry shapes accuracy, which we evaluate in Sec.4 on real data: deviations from these predictions quantify the role of correlations in the classification task.

Assumption 3.1 (*uncorrelated statistics*). For each class $\mu = 1, \dots, P$, the centroid is drawn as $\mathbf{c}_\mu \sim_{\text{iid}} \mathcal{N}(\mathbf{0}, I_N)$. Conditional on the centroids, examples of class k are generated by sampling $\mathbf{x} = \mathbf{c}_k + \delta\mathbf{x}$, with residuals drawn as $\delta\mathbf{x} \sim_{\text{iid}} \mathcal{N}(\mathbf{0}, R^2 I_N)$.

Assumption 3.2 (*high-dimensional geometry*). We work in the high-dimensional limit $N \rightarrow \infty$, with $P = N^{\mathcal{O}(1)}$ classes and a radius $R = R(N)$ that may scale with N .

We refer to Assumptions 3.1-3.2 jointly as the isotropic model.

Theorem 3.3 (*isotropic model accuracy*). Under the isotropic model,

$$\text{Acc}^{\text{iso.}}(R, N, P) = \mathbb{E}_{z \sim \mathcal{N}(0,1)} \left[\Phi^{P-1} \left(\frac{Rz + \sqrt{N}}{\sqrt{R^2 + 1}} \right) \right], \quad (3)$$

where Φ is the standard Gaussian CDF.

The supporting lemmas (centroid-overlap CLT, self-averaging of Acc, and the joint Gaussian distribution of the logits), together with the proof of Theorem 3.3, are given in Appendix C.1.

Predictions. Several consequences follow directly from Theorem 3.3:

- (i) For any $R = \mathcal{O}(1)$, the classes are well separated, and classification is trivial: $\text{Acc}^{\text{iso.}} \rightarrow 1$.
- (ii) In the infinite radii limit, $R/\sqrt{N} \rightarrow \infty$, the isotropic residual dominates the input, which thus retains no information about the true class: all P classes are equally likely to be predicted, and accuracy collapses to chance, $\text{Acc}^{\text{iso.}} \rightarrow 1/P$.
- (iii) Non-perfect accuracy requires large radii, $\text{Acc}^{\text{iso.}} < 1 \Rightarrow R = \Omega(\sqrt{N/\ln P})$.
- (iv) In the high-dimensional regime, the accuracy is self-averaging: its variance across centroid realizations is $\mathcal{O}((PR^2)^{-1})$. We note that according to (iii), R also scales with N for nontrivial accuracy.

These predictions are rigorously derived in Appendix C.1 (Corollaries C.1.6-C.1.9), and verified numerically with synthetic data in Appendix Fig.6.

4 Empirical Results

In this section we characterize the geometric structure of trained representations that governs prototype classification. We begin from the predictions of the isotropic model (Sec.3) and show that they are

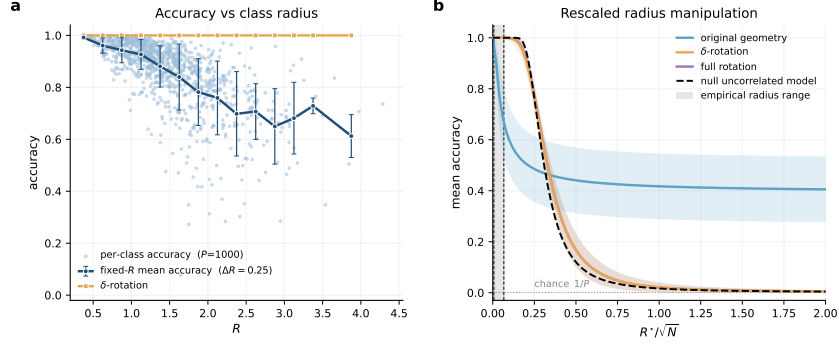


Figure 2: **Geometric signatures of multiclass classification** (a) Per-class accuracy vs. empirical radius R : sharp decrease across the narrow empirical R range and a wide spread at fixed R . After random rotation of within-class variability $\delta\mathbf{x}$ (δ -rotation), the accuracy is perfect. (b) We artificially scale the variability $\delta\mathbf{x} \rightarrow (R^*/R)\delta\mathbf{x}$. The accuracy decreases fast but stabilizes at 39% accuracy for $R^* \rightarrow \infty$ due to strong true centroid correlations (blue line). After random rotation of the variability δ -rotation, the accuracy is well described by the isotropic model ($\text{MSE} = 1.26 \times 10^{-3}$, orange line). Additional breaking of centroid-centroid correlations (full rotation) has a very small effect; lines overlap.

not satisfied by real representations. This pattern of failure is informative: it identifies which correlations dominate the geometry responsible for accuracy, and isolates the components needed to build the analytical framework of Sec.5. Results are presented for DINOv3 7B [Siméoni et al., 2025] representations on ImageNet-1K [Deng et al., 2009], and later replicated in many models and modalities (Sec. 5.3).

Accuracy changes on an $\mathcal{O}(1)$ radius scale: The empirical class radii lie in a narrow $R \in [0.4, 4.3]$ range. In Sec.3, we saw that an uncorrelated model predicts that accuracy should remain unchanged throughout $\mathcal{O}(1)$ radii, and decline only in $R = \Omega(\sqrt{N/\ln P}) \sim 20$. However, we observe a dramatic change in accuracy, ranging from 100% to 61% averaged over a small R window (see Fig.2a). The sharp accuracy decline therefore indicates structure that is not described by the isotropic model.

Large class-to-class variability at fixed radius: The isotropic model predicts that class-to-class fluctuations are of order $1/R\sqrt{P}$ ($\sim 1\% - 3\%$). Empirically, however, the spread of accuracies at fixed R is large. For example, the 89 classes in the range $R \in [1.2, 1.3]$ have a mean accuracy of 90.4%, with a minimum accuracy of 61.8%, and a maximum of 99.6% (see Fig.2a). The radius alone therefore cannot explain classification performance: additional class-dependent geometric structure must determine which classes are easy to classify and which are difficult.

4.1 Geometric manipulations

To causally test which components cause the deviation from the isotropic model, we use geometric manipulations to probe specific structures of the representations. We apply two geometric manipulations:

- (i) **Variability inflation**, where we set each class radius to a chosen R^* artificially by scaling its deviations, $\mathbf{x}(R^*) = \mathbf{c}_k + (R^*/R)\delta\mathbf{x}$, where R is the empirical radius as defined in Sec.2.
- (ii) **Geometric rotations**, where we surgically destroy specific correlations in the data while preserving all other aspects of the representation geometry.

Strong residual-true-centroid correlations: We first examine the representations under variability inflation $\mathbf{x}(R^*) = \mathbf{c}_k + (R^*/R)\delta\mathbf{x}$. Naively, as $R^* \rightarrow \infty$ one expects chance accuracy, as we saw in the isotropic model. Empirically, the large-radius accuracy instead stays far above chance - accuracy approaches 39.2%, over two orders of magnitude above $1/P = 0.1\%$ (Fig.2b). The within-class variability is thus far more correlated with the true centroid than with the rival centroids. To quantify this directly, we measure the projection variance along each centroid direction σ_μ (see Sec. 2). The true-class σ_k exceeds all rivals in 93% of classes. Moreover, this dominance is itself predictive of a well-structured class geometry: in the 930 classes where σ_k exceeds all rivals, the mean accuracy is 87.4%, whereas in the remaining 70 classes it drops to 60.1%.

Residual-centroid correlations are the main drivers of errors: We manipulate the representations by rotating the residuals $\delta\mathbf{x}$ around their own class centroid (δ -rotations): for each class k we replace $\mathbf{x} = \mathbf{c}_k + \delta\mathbf{x}$ by $\mathbf{x}^{\delta\text{rot}} = \mathbf{c}_k + O_k\delta\mathbf{x}$, with O_k a class-specific random orthogonal matrix (see Fig. 1b for illustration). This preserves the centroid-centroid correlations, the class radii, and the example-example relations, while breaking the correlations between $\delta\mathbf{x}$ and the centroid directions (residual-centroid correlations). At the empirical radius, the class-averaged accuracy rises from the baseline of 85.5% to 100%, completely eliminating classification errors (see Fig.2a, orange line). This is achieved *without* breaking centroid-centroid correlations.

Breaking residual-centroid correlations recovers the isotropic model: To look for any further role of residual-centroid correlations, we go outside the empirical radius range, where δ -rotation has already saturated the accuracy. We set each class radius to a chosen R^* by variability inflation. Strikingly, after removing only the residual-centroid correlations, the data is already well described by the isotropic model accuracy (Eq.3), with $\text{MSE} = 1.26 \times 10^{-3}$. The accuracy declines only at $R = \Omega(\sqrt{N/\ln P})$, far beyond the empirical R range (see Fig.2b). The accuracy at $R^* \rightarrow \infty$ is chance (0.1%), and the per-class standard deviation around the mean accuracy shrinks by an order of magnitude to 2.3×10^{-3} compared to 1.7×10^{-2} in the unmanipulated representations.

Centroid-centroid correlations play a minor role: To test the role of centroid-centroid correlations, we apply a stronger manipulation: rotation of the whole class manifold around the origin by another random orthogonal matrix after δ -rotations, $\mathbf{x}^{\text{full-rot}} = U_k(\mathbf{c}_k + O_k\delta\mathbf{x})$, which breaks centroid-centroid correlations on top of the δ -rotation. Even at these larger radii the extra effect is small: the full-rotation curve differs from the δ -rotation curve by only $\text{MSE} = 2.0 \times 10^{-5}$ (see Fig. 2b). We infer that centroid-centroid correlations play a minor role in accuracy. This echoes the picture from neural collapse [Papayan et al., 2020], where the centroids are as spread as possible.

Together, these empirical results identify the key geometric structure of trained representations: the class-relevant geometry is governed by strong correlations between within-class variability and a sparse set of centroid directions—dominated by the true-class centroid.

5 Centroid-Aligned Variability Model

The empirical results of Sec. 4 identify the geometry of prototype classification in pretrained representations: Within-class variability $\delta\mathbf{x}$ is *centroid-aligned*. Surgically removing only the $\delta\mathbf{x}$ -centroid correlations drives accuracy to $\sim 100\%$, while breaking centroid-centroid correlations has only a secondary effect.

We introduce the centroid-aligned model that incorporates this geometry. The resulting theory takes as input geometric quantities measurable from the representation, together with a single global renormalization of the class radius (Sec. 5.2), and predicts per-class accuracy across architectures and modalities. All proofs, lemmas, and intermediate computations are deferred to Appendix C.2; the main text retains only the core assumptions and central results. We stress that the following assumptions are introduced to define an analytically solvable model and are not intended as exact claims about the statistics of the data. Nevertheless, the model is motivated by empirical observations (Sec. 4) and, as we show below, yields accurate predictions across models spanning different architectures and modalities. We therefore believe that it captures important aspects of real model representations.

Definition 5.1 (rival support). Fix a class k and an integer K . Define its rival support $\mathcal{R}_K$ as the set of indices $\mu \neq k$ corresponding to the K largest values of σ_μ (Def. 2.2), and set $\mathcal{R}_K^+ \equiv \{k\} \cup \mathcal{R}_K$. For each $\mu \in \mathcal{R}_K$, define the true-rival centroid correlation by $g_\mu = \hat{\mathbf{c}}_k \cdot \hat{\mathbf{c}}_\mu$.

Definition 5.2 (standardized centroid projections). Conditioned on class k , each example is represented by the following projection coefficients:

$$s_\mu = \frac{\delta\mathbf{x} \cdot \hat{\mathbf{c}}_\mu}{\text{std}_k[\delta\mathbf{x} \cdot \hat{\mathbf{c}}_\mu]}, \quad \mu \in \mathcal{R}_K^+ \quad (4)$$

Assumption 5.3 (centroid-aligned variability model). Let $G_{\mu\nu} = \hat{\mathbf{c}}_\mu \cdot \hat{\mathbf{c}}_\nu$ be the centroids Gram matrix. Assume that the restricted centroid Gram matrix of the rival support $(G_{\mu\nu})_{\mu,\nu \in \mathcal{R}_K^+}$ is invertible. We assume the following generative model of examples of class k

$$\mathbf{x}(s) = \|\mathbf{c}_k\| \left[\hat{\mathbf{c}}_k + R \left(\sum_{\mu,\nu \in \mathcal{R}_K^+} \sigma_\mu s_\mu G_{\mu\nu}^{-1} \hat{\mathbf{c}}_\nu + \sigma_\perp s_\perp \hat{\mathbf{e}}_\perp \right) \right] \quad (5)$$

where $\hat{\mathbf{e}}_{\perp}$ is perpendicular to the $\text{span}(\hat{\mathbf{c}}_{\mu \in \mathcal{R}_k^+})$, σ_{μ}^2 is the variance fraction along the centroid μ direction, and R is the empirical class radius (see Def. 2.2). Eq. 5 encodes the geometry identified in Sec. 4 by placing the residuals $\delta \mathbf{x}$ along rival centroid directions in the rival support.

Assumption 5.4 (Gaussian statistics). We assume the projection coefficients s_{μ} are jointly Gaussian with zero mean and unit variance, and that their covariance has the following form

$$\text{Cov}(s_{\mu}, s_{\nu}) = \delta_{\mu\nu} (1 - \rho_{\mu}^2) + \rho_{\mu} \rho_{\nu}, \quad (6)$$

where $\rho_{\mu} = \mathbb{E}_k [s_k s_{\mu}]$, and by definition $\rho_k = 1$. These are the Pearson correlations between the residual projections $\delta \mathbf{x} \cdot \hat{\mathbf{c}}_k$ and $\delta \mathbf{x} \cdot \hat{\mathbf{c}}_{\mu}$. Equivalently, all dependence among the rival projections is captured by s_k : conditioned on s_k , they are mutually independent.

Theorem 5.5 (centroid-aligned model accuracy). Under assumptions 5.3-5.4

$$\text{Acc}_k^{\text{th}}(R) = \mathbb{E}_{z \sim \mathcal{N}(0,1)} \left[\prod_{\mu \in \mathcal{R}_K} \Phi \left(\frac{\sigma_k - \rho_{\mu} \sigma_{\mu}}{\sigma_{\mu} \sqrt{1 - \rho_{\mu}^2}} z + \frac{1 - g_{\mu}}{R \sigma_{\mu} \sqrt{1 - \rho_{\mu}^2}} \right) \right]. \quad (7)$$

The inputs to Eq. 7 are entirely geometric and class-specific: the empirical radius R , the rival projected variances $\{\sigma_{\mu}\}$, the centroid-centroid overlaps $\{g_{\mu}\}$, and the correlation terms ρ_{μ} are all measured directly from the representations. For convenience, we summarize our notations and how we measure them in Table. 2.

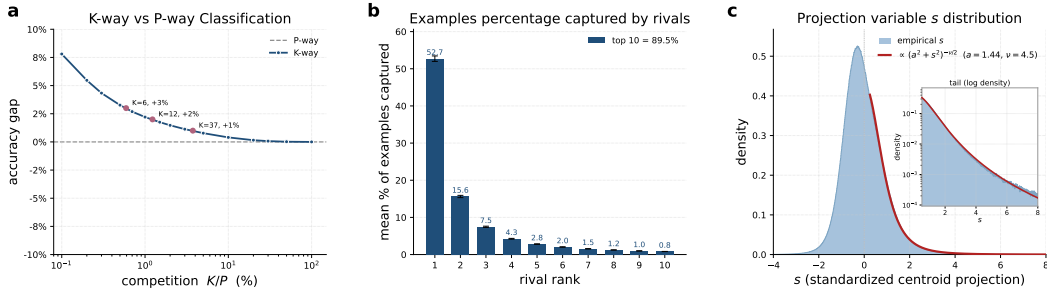


Figure 3: Sparsity and heavy tail statistics: (a) K-way classification of rivals chosen by centroid projection variance σ_{μ} converges quickly to the full P -way classification. (b) We define the strongest rival of each example $\mu^*(\mathbf{x}) = \arg \max_{\mu \neq k} t_{\mu}(\mathbf{x})$, and count the number of examples captured by each rival. The strongest rival of each class captures more than half of the examples on average, with 90% of the examples captured by only 14.4 rivals. (c) The standardized centroid projections s_{μ} (see Eq.4) have heavy tail power law statistics, which inflate the extreme values that control classification accuracy. The magnified panel shows the log density of the far tail.

5.1 Sparsity

The rival count K is a hyperparameter of the theoretical model controlling how many rival directions enter the prediction in Eq. 7. Because Φ approaches one quickly as its argument grows, small projections σ_{μ} have very little effect on the overall accuracy. Thus, when ranking rivals by their projection magnitudes σ_{μ} , predictive power plateaus rapidly once K exceeds the head of the spectrum.

We assess the effective number of rivals both theoretically and empirically. We first replace the full P -way task with a K -way task restricted to the top rivals ranked by σ_{μ} . We find that only 6 rivals are needed to come within 3% of the full P -way classification accuracy, 12 within 2%, and 37 within 1% (see Fig.3a). The competition is thus not about the full set of P (1000) classes, but rather a sparse set of strongly correlated rivals. In Appendix E, we compare several rival-selection methods and show that ranking rivals by σ_{μ} yields results comparable to more complex selection rules, while remaining simple, interpretable, and easy to compute. In particular, it outperforms selection based on centroid proximity, providing additional evidence that centroid correlations play only a secondary role in the task.

We further examine whether the correlations are homogeneously distributed across the rivals or dominated by a few. For each example $\mathbf{x}$ of class k , we record its strongest rival, $\mu^*(\mathbf{x}) = \arg \max_{\mu \neq k} t_{\mu}(\mathbf{x})$, and count the number of distinct rivals that appear across the ~ 1300 examples

model	class	top-5 rivals (by σ_μ)
Qwen3-32B	<i>happy</i>	excite, proud, glad, surprised, curious
	<i>bad</i>	sad, sorry, scare, angry, tired
	<i>voice</i>	loud, noise, sound, big, funny
	<i>smile</i>	reply, nod, say, think, tell
	<i>cake</i>	delicious, cookie, food, ice, big
	<i>forest</i>	park, garden, grass, world, sky
DINOv3 ViT-7B	<i>ground</i>	floor, grass, swing, slide, garden
	<i>leopard</i>	snow leopard, jaguar, cheetah, impala, tusker
	<i>king crab</i>	Dungeness crab, rock crab, hermit crab, plate, American lobster
	<i>flute</i>	panpipe, oboe, ocarina, bassoon, stage
	<i>orange</i>	lemon, banana, grocery store, Granny Smith, strawberry

Table 1: Rival classes selected by the projection’s variance σ_μ are semantically related words (Qwen3-32B, TinyStories) or visually similar and co-occurring objects (DINOv3 ViT-7B, ImageNet). For more examples see also Fig. 1a.

of that class. Competition is highly concentrated: on average only 67 distinct rivals appear, 14.4 of them cover 90% of the examples, and one rival alone accounts for more than half (Fig. 3b).

The theory’s predictions also saturate quickly as the number of rivals K increases. We find that $K \in [10, 20]$ is already enough in most models to achieve accurate predictions, with very small gain from increasing K further (see Appendix E). The results in the main text are shown for $K = 20$ for all models.

We also note that in many cases, the rivals chosen by σ_μ carry semantic meaning. In the Qwen3-32B [Yang et al., 2025] model on TinyStories [Eldan and Li, 2023], many rival words are similar in both grammatical role ("smile" $\mapsto$ "reply", "nod", "say") and semantic meaning ("forest" $\mapsto$ "park", "garden", "grass"). In vision (DINOv3 ViT-7B), we observe either similar objects ("leopard" $\mapsto$ "jaguar", "cheetah") or frequently co-occurring objects ("king crab" $\mapsto$ "plate", "flute" $\mapsto$ "stage"). For more examples, see Table 1 and Fig. 1a.

5.2 Non-Gaussian tails and a global radius renormalization

The theoretical accuracy predicted by Eq. 7 alone consistently overestimates the accuracy on all the models we checked (see Appendix F). Our main simplifying assumption (Assumption 5.4) is that the statistics of the standardized centroid projections s_μ are Gaussian. Empirically, the projections have markedly heavy-tailed statistics (see Fig. 3c). Although the variance of s_μ is normalized, classification depends on extreme statistics ($\max_\mu t_\mu$), and heavier tails inflate these extremes relative to a Gaussian of equal variance. We find that the right tail of the variables s_μ that controls the max projection on a certain centroid is well fitted by a Student’s-t power-law distribution $p(s) \propto (a^2 + s^2)^{-\nu/2}$ (see Fig. 3c), where the exponent ν shapes the decay of the tail.

We adopt the simplest solution: a uniform global variance inflation of all manifolds $R \mapsto \lambda R$ in Theorem 5.5, which adjusts the scale of the Gaussian theory such that its extreme matches the empirical heavy-tailed distribution.

The global parameter λ is fitted once per model by

$$\lambda^* = \arg \min_{\lambda} \left(\sum_{k=1}^P \left[\text{Acc}_k^{\text{emp}}(R) - \text{Acc}_k^{\text{th}}(\lambda R) \right]^2 \right), \quad (8)$$

Notation	Name	Definition
$\mathbf{c}_k$	Centroid	$\mathbb{E}_k[\mathbf{x}]$
$\hat{\mathbf{c}}_\mu$	Unit centroid	$\mathbf{c}_\mu / \ \mathbf{c}_\mu\ $
$\delta \mathbf{x}$	Residual	$\mathbf{x} - \mathbf{c}_k$
R^2	Radius ²	$\mathbb{E}_k[\ \delta \mathbf{x}\ ^2] / \ \mathbf{c}_k\ ^2$
σ_μ^2	Variance fraction	$\mathbb{E}_k[(\delta \mathbf{x} \cdot \hat{\mathbf{c}}_\mu)^2] / \mathbb{E}_k[\ \delta \mathbf{x}\ ^2]$
s_μ	Std. projection	$(\delta \mathbf{x} \cdot \hat{\mathbf{c}}_\mu) / \text{std}_k[\delta \mathbf{x} \cdot \hat{\mathbf{c}}_\mu]$
ρ_μ	Proj. correlation	$\text{Corr}_k(\delta \mathbf{x} \cdot \hat{\mathbf{c}}_\mu, \delta \mathbf{x} \cdot \hat{\mathbf{c}}_\mu)$
g_μ	Centroid overlap	$\hat{\mathbf{c}}_k \cdot \hat{\mathbf{c}}_\mu$
$\mathcal{R}_K$	Rival set	$\{\mu \neq k : K \text{ largest } \sigma_\mu\}$

Table 2: Definitions. $\mathbb{E}_k$ is defined as average over all samples $\mathbf{x}$ belong to the manifold k

Thus λ is a single global calibration of non-Gaussian extreme-value statistics of a specific model, not a per-class fitting parameter. Empirically λ^* is well correlated with the inverse tail exponent $1/\nu$ of the model, $r = 0.90$ (see Fig. 4e). This relation is global across 36 models spanning vision, audio, and language, supporting our interpretation of λ as a compensation for non-Gaussian statistics.

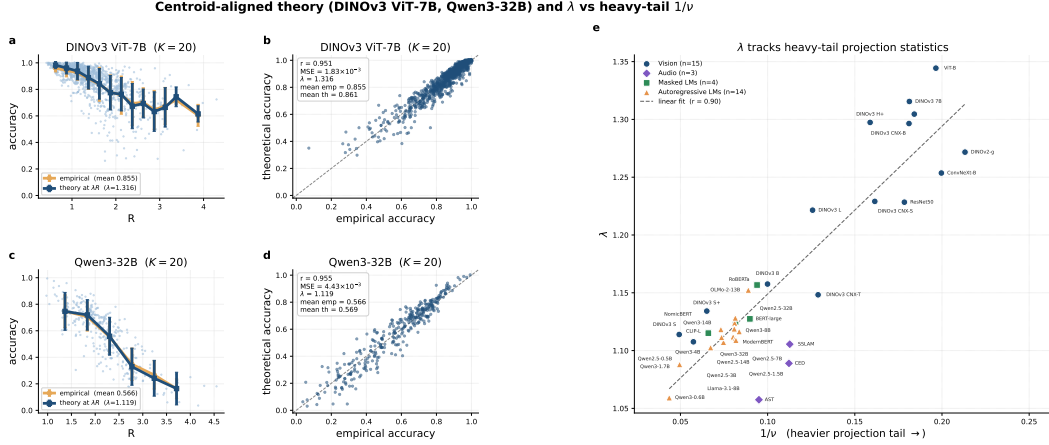


Figure 4: **Centroid-aligned theory matches per-class accuracy across modalities.** (a, c) Per-class accuracy vs. empirical radius R for DINOv3 ViT-7B on ImageNet ($P=1000$, $K=20$) and Qwen3-32B on TinyStories next-token prediction ($P=320$, $K=20$). Per-class scatter is presented with empirical binned mean. The λ -rescaled theory ($\lambda^*=1.316$, 1.119 resp.) accounts for the empirical accuracy accurately. (b, d) Accurate per-class prediction of the accuracy. $MSE=1.83 \times 10^{-3}$ (vision) and 4.43×10^{-3} (language), Pearson $r=0.951$ and 0.955 . (e) Optimal λ^* vs. heavy tail measurement ($1/\nu$) on representations for 15 vision backbones, 14 autoregressive language models, 4 masked language models, and 3 audio models. Overall Pearson correlations between λ and $1/\nu$ are $r=0.90$, confirming that λ is a byproduct of the heavy-tail shape of the projection variables.

5.3 Shared geometry across architectures and modalities

The empirical structure of Sections 4-5 was characterized on DINOv3 ViT-7B. Table 3 reports the same measurements across a diverse set of vision, audio, and language backbones: Every observation persists: R is in an $\mathcal{O}(1)$ range for any N , with large accuracy fluctuations for fixed R (independent of N). δ -rotation drives the class-averaged accuracy to near-perfect on all backbones; the saturation accuracy ($R^* \rightarrow \infty$) is consistently orders of magnitude above chance, indicating true-class dominance among the residual correlations. $K_{90\%}$ is the number of rivals that are the strongest competitor for 90% of a class’s examples. This rival set is sparse, $K_{90\%} \ll P$ in all cases. The projection s_μ right tail is heavy on every backbone. The geometric characterizations required by the theory of Section 5 are therefore present and consistent across different architectures, modalities, and training procedures. Appendix H and Figs. 10-13 report the theory-empirical comparison for all 36 backbones.

6 Geometry and Accuracy Across Model Scales

The theory of Sec. 5 reduces accuracy to a handful of ingredients of the centroid-aligned geometry. We use these quantities as a diagnostic lens across the DINOv3 and Qwen3 model families at different scales ($S(21M)/S_+(29M)/B(86M)/L(300M)/H_+(840M)/7B$, $0.6B/1.7B/4B/8B/14B/32B$ resp.). The theory reveals which geometric reorganizations drive the improvements in classification performance of larger models.

A natural concern is that the accuracy gains of larger models are a trivial variability-reduction effect — that bigger models simply produce tighter manifolds. The Qwen3 family rules this out directly: the mean radius actually grows with scale ($\langle R \rangle = 1.98 \rightarrow 2.20$ from $0.6B$ to $32B$), yet accuracy rises monotonically from 41% to 57%, the opposite of what a naive approach predicts. In the DINOv3 family, the radius does shrink with scale, but the trend breaks at the largest model: ViT-7B carries a mean radius 0.21 larger than ViT-H⁺ yet matches its accuracy. Reduced variability, therefore, cannot be the mechanism behind improved generalization. Our other geometric measures are independent of

Table 3: Universality of the centroid-aligned geometry. For each backbone, we report: the number of classes P ; feature dimension N ; per-class radius range; per-class mean accuracy $\langle \text{Acc} \rangle$ and its fluctuation around the mean accuracy at a fixed R ; mean accuracy after δ -rotation; saturation accuracy at $R^* \rightarrow \infty$ (controlled only by the residual $\delta \mathbf{x}$); $K_{90\%}$ (mean rivals covering 90% of examples); right-tail Student- t tail index ν ($K=20$).

Backbone	P	N	R range	$\langle \text{Acc} \rangle$	Fixed R fluctuations	δ -rot	$R^* \rightarrow \infty$	$\langle K_{90\%} \rangle$	ν
<i>Vision – ImageNet-1K</i>									
DINOv3 ViT-7B/16	1000	4096	[0.4, 4.3]	0.855	0.099	1.000	0.392	14.4	4.5
DINOv3 ConvNeXt-L	1000	1536	[0.2, 3.1]	0.814	0.100	1.000	0.312	14.2	4.4
CLIP ViT-L/14	1000	768	[0.6, 2.7]	0.792	0.124	0.999	0.077	13.9	16.4
ViT-B/16 (IN21K)	1000	768	[0.5, 2.7]	0.853	0.088	1.000	0.265	14.1	4.1
<i>Audio – VGGSound</i>									
SSLAM	310	768	[0.1, 3.7]	0.628	0.155	0.999	0.067	11.0	7.9
CED	310	768	[0.1, 4.0]	0.628	0.160	1.000	0.062	11.3	7.9
<i>Language - TinyStories masked-token</i>									
RoBERTa-large	320	1024	[0.6, 2.6]	0.767	0.108	1.000	0.120	13.3	9.6
ModernBERT-large	320	1024	[0.6, 2.7]	0.734	0.112	1.000	0.118	12.4	11.3
BERT-large	320	1024	[0.6, 2.7]	0.710	0.112	1.000	0.120	15.8	10.1
<i>Language - TinyStories next-token</i>									
Qwen3-8B (L34, -2)	320	4096	[0.9, 4.2]	0.540	0.118	1.000	0.107	33.7	11.3
OLMo2-13B (L36, -4)	320	5120	[1.0, 4.8]	0.563	0.116	1.000	0.112	35.8	10.2
Llama-3.1-8B (L29, -3)	320	4096	[0.8, 4.1]	0.550	0.119	1.000	0.101	31.7	12.4

R , so they isolate genuine reorganizations of the geometry across scale, and correctly identify the components that make the larger models’ representations better.

We specifically look at the worst rival μ^* , measured by the maximum σ_μ (most correlated centroid among the variability) for each class, and average its geometric statistics over all classes. The theory exposes a coherent set of reorganizations, shared among the two families. Within-class variability becomes more aligned with the true centroid and less aligned with the rival (the ratio σ_k/σ_μ^* increases, σ_μ^* decreases in both families). The centroids decorrelate - the overlap between the true-centroid and its hardest rival g_μ^* drops - and the projections of the residual on the centroids also become less correlated (ρ^* decreases). In language, we see similar trends to vision but with less favorable values, which explains the overall lower accuracy. Finally, representations become more heavy-tailed with scale (ν decreases, $19.3 \mapsto 4.5$ in vision and $21.9 \mapsto 11.4$ in language). While we cannot directly link this trend to improved accuracy, it is consistent.

A large body of literature [Peters et al., 2018, Liu et al., 2019a, Tenney et al., 2019, Skean et al., 2025] reports that in many cases, the best representations for tasks are found in the middle hidden layers, and not in the final layer. We add our geometric angle to it in finding that in autoregressive models, the best prototype classification is consistently 1-3 layers before the final layer. We report here the geometry for the best layer found for prototype classification, and analyze how the geometry changes across depth in Appendix G.

7 Discussion

This work identifies geometric organization principles for multiclass representations: prototype accuracy is controlled by a sparse, centroid-aligned code embedded in the full activation space. Across vision, audio, and language models, and across architectures and training objectives, the same geometry recurs. Within-class variability is not isotropic around each class center; its classifier-relevant components are arranged along the true centroid and a small set of rival centroids. The evidence is causal as well as descriptive: breaking this alignment, while preserving the remaining parts of the geometry, nearly eliminates prototype errors. Thus the full N -dimensional representation can be replaced, for the purpose of predicting class accuracy, by a much smaller set of task-aligned coordinates.

The theory makes this reduction quantitative. Per-class accuracy is predicted from the geometric measures of centroid-projection statistics, and a single global scale λ that captures the effect of heavy tails in the marginal projection distributions. Radius alone is insufficient to explain the results. We show that as models become larger, the improved accuracy is not due to reduced class variance

Table 4: **Geometric measures across model scales:** DINOv3 ViT family on ImageNet-1K (top block) and the Qwen3 language-model family on TinyStories (bottom block). The geometric quantities are defined in Sec. 5 and Table 2; the superscript * flags the hardest competitor of a class, identified by σ_μ . Best measure within each family marked in bold.

Model	$\langle R \rangle$	σ_k/σ_μ^*	σ_μ^*	g_μ^*	ρ_μ^*	ν	Acc
DINOv3 ViT-S	1.68	1.06	0.172	0.574	0.558	19.3	0.706
DINOv3 ViT-S+	1.60	1.09	0.172	0.555	0.535	14.4	0.738
DINOv3 ViT-B	1.52	1.19	0.163	0.542	0.529	9.0	0.794
DINOv3 ViT-L	1.39	1.36	0.147	0.447	0.322	7.0	0.845
DINOv3 ViT-H+	1.29	1.59	0.164	0.236	-0.011	5.3	0.855
DINOv3 ViT-7B	1.50	2.01	0.120	0.289	0.038	4.5	0.855
Qwen3-0.6B (L27, -1)	1.98	0.85	0.241	0.576	0.515	21.9	0.406
Qwen3-1.7B (L26, -2)	2.09	0.95	0.212	0.542	0.494	13.9	0.474
Qwen3-4B (L34, -2)	2.05	0.98	0.202	0.526	0.463	12.7	0.519
Qwen3-8B (L34, -2)	2.12	0.99	0.188	0.514	0.453	11.3	0.540
Qwen3-14B (L37, -3)	2.10	1.00	0.186	0.504	0.447	11.3	0.556
Qwen3-32B (L61, -3)	2.20	1.01	0.175	0.503	0.451	11.4	0.566

but because of reorganization of the geometry: variability shifts toward the true-class coordinate, correlations with the strongest rivals decrease, and centroid directions become more decorrelated.

Importantly, the theory tracks not only prototype classification but also standard linear classification (Appendix A.1). For linear readouts, a train–test split is essential, since a readout can fit arbitrary labels whenever the data remain separable. Prototype classification, by contrast, achieves high accuracy only for genuinely well-organized geometry — shuffled labels or random representations perform near chance — so the train–test split (calculating the centroids on different data than the one classified) contributes only finite-sample effects, negligible for large sample sizes when train and test are drawn from the same distribution.

This perspective complements existing theories of representational geometry. Manifold-capacity theory [Chung et al., 2018, Cohen et al., 2020] models class manifolds as ellipsoids and characterizes their linear separability under random binary dichotomies. In that setting, the primary component of the geometry is the random structure of the whole P manifolds, and correlations enter as corrections on top of the random theory [Wakhloo et al., 2023]. We showed that in our task the centroid-aligned component of within-class variability is the dominant part, while the random fluctuations of inter-centroid overlaps have little to no effect. The two frameworks are complementary regimes of the same underlying geometry, and it would be interesting to further investigate where the two descriptions meet.

The sparse code uncovered here also suggests links to interpretability and neuroscience. Sparse autoencoders and dictionary-learning methods seek low-dimensional, semantically meaningful directions in activations [Makhzani and Frey, 2013, Huben et al., 2024, Templeton et al., 2026, Gao et al., 2025]; our results identify a complementary sparse class-supervised set of directions whose functional role is directly tied to classification. In parallel, deep networks have been shown to predict responses in primate ventral-stream areas, including IT cortex [Yamins et al., 2014, Schrimpf et al., 2020], and object-manifold theory has been applied to biological populations [Froudarakis et al., 2020, Sorscher et al., 2022, Kuoch et al., 2024]. It would be interesting to examine whether the universal geometry we identify across many artificial models also applies to representations in the brain.

Finally, the theory points to practical applications. Because its ingredients are compact and inexpensive to estimate, they can serve as diagnostics for representation quality, to select which pretrained model is most suitable for fine-tuning on a new task [Kornblith et al., 2018]. They could also suggest geometry-aware objectives: training could explicitly encourage strong own-centroid alignment, weak rival alignment, and sparse competition, while distillation could match the teacher and student geometry rather than the entire N -dimensional representation or the logit distribution [Hinton et al., 2015]. Following these directions would turn the centroid-aligned description from a predictive theory of existing representations into a constructive principle for designing better ones.

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A Prototype Classification

A.1 Prototype classification and linear readout comparison

One natural concern regarding our work is that it is applied only to prototype classification, with no relation to the more common linear readout. We test this by splitting the data into 90% train and 10% validation, training a max-margin SVM (using the Crammer–Singer method [Crammer and Singer, 2001]) on the train data, and evaluating on the validation set. We compare it with computing the centroids over the 90% train representations, and computing the prototype scores on the 10%. The gap between the two methods is small (see Fig.5 and Table 5), especially in vision (0.047) and autoregressive models (0.063). We note that in the bidirectional language models, the gap is substantially larger than in the other modalities (0.121), which is also reflected in a lower per-class correlation between the two classifiers ($r = 0.58$, against 0.87 in vision and 0.76 in the autoregressive models). The additional structure that allows the linear readout to separate the data better than the centroids in these models remains an open question.



Figure 5: Centroid vs trained linear-readout accuracy, per backbone. The gap between methods is small (see also Table 5).

B Experimental Setup

The main text evaluates frozen representations from pretrained backbones spanning vision, language, and audio, across transformer and convolutional architectures, supervised, self-supervised, and contrastive objectives. This appendix gives the full list of datasets, models, and extraction details. No backbone is fine-tuned or changed for our analysis.

B.1 Datasets

Vision-ImageNet-1K: We use ImageNet-1K (ILSVRC 2012) [Deng et al., 2009], the standard 1000-class image benchmark with $\approx 1.28\text{M}$ images.

Backbone	r	ΔAcc
<i>Vision (ImageNet-1K)</i>		
ViT-B/16 (IN21K)	0.937	-0.005
ConvNeXt-B (IN1K)	0.893	-0.001
DINOv3 ViT-L/16	0.885	+0.028
DINOv3 ViT-H ⁺ /16	0.873	+0.028
DINOv3 ConvNeXt-S	0.872	+0.058
CLIP ViT-L/14	0.869	+0.057
DINOv3 ConvNeXt-T	0.865	+0.076
DINOv3 ViT-S ⁺ /16	0.864	+0.084
ResNet-50	0.861	+0.016
DINOv3 ViT-S/16	0.859	+0.100
DINOv3 ViT-B/16	0.857	+0.058
DINOv3 ConvNeXt-B	0.856	+0.054
DINOv3 ConvNeXt-L	0.854	+0.052
DINOv3 ViT-7B/16	0.839	+0.030
DINOv2 ViT-g/14	0.791	+0.074
<i>mean</i>	0.865	+0.047
<i>Audio (VGGSound)</i>		
SSLAM	0.821	+0.084
AST	0.815	+0.091
CED	0.808	+0.088
<i>mean</i>	0.815	+0.088

Backbone	r	ΔAcc
<i>Masked LM (TinyStories, masked-word)</i>		
ModernBERT-large	0.584	+0.126
NomicBERT	0.581	+0.159
BERT-large	0.568	+0.102
RoBERTa-large	0.565	+0.096
<i>mean</i>	0.575	+0.121
<i>Autoregressive LM (TinyStories, next-token)</i>		
Qwen3-4B	0.785	+0.070
Qwen3-1.7B	0.777	+0.089
Qwen3-8B	0.772	+0.050
Qwen2.5-0.5B	0.769	+0.102
Qwen3-0.6B	0.765	+0.122
Llama-3.1-8B	0.764	+0.041
Qwen2.5-3B	0.762	+0.076
Qwen3-32B	0.761	+0.019
Qwen3-14B	0.754	+0.038
OLMo-2-13B	0.750	+0.029
Qwen2.5-1.5B	0.749	+0.102
Qwen2.5-32B	0.747	+0.036
Qwen2.5-7B	0.740	+0.052
Qwen2.5-14B	0.736	+0.050
<i>mean</i>	0.759	+0.063
Overall	0.787	+0.065

Table 5: Per-class Pearson correlation r between trained linear-readout and prototype-classifier accuracy, together with the mean accuracy gap ΔAcc (readout minus prototype) on the held-out split. The two classifiers align closely, with the exception of masked language models (see Sec. A.1).

Language-TinyStories: For language, the task is token prediction on the validation dataset of TinyStories [Eldan and Li, 2023], a corpus of 20K short synthetic children’s stories. We treat token prediction as a multi-class classification problem.

Content words and lemmatizing: In order to get semantically meaningful representations, we keep only data points of *content words*. Following [Arps et al., 2024], content words are those whose part-of-speech is an open lexical class of the Universal Dependencies scheme [De Marneffe et al., 2021]: NOUN, PROPN, VERB, ADJ, or ADV. We assign part-of-speech tags and lemmas with the spaCy pipeline [Honnibal et al., 2020], and unify all tokens sharing a lemma into a single class (for example, "say, said, says" or "friend, friends"). We keep only classes with at least 1000 occurrences, so that the manifold geometry is well estimated. Finally, to obtain an identical class set across models with different tokenizers, we discard the 11 lemmas that are not realized consistently in every tokenizer - fragmented or unified into/from sub-words differently in at least one model ("benny, bunny, fire, frog, grandma, grow, mia, mommy, sweet, window, yummy", for example one tokenizer treats "fireman" $\mapsto$ "fire+man"). This leaves 320 classes shared by all backbones.

Masked-word task (bidirectional encoders). We mask a single content-word position at a time, run a full bidirectional forward pass, and read out the encoder’s representation at the masked position. The representation is the final hidden state feeding the model’s masked-LM head. The class label is the *lemma* of the masked word, restricted to the same open-class content-word construction as above.

Next-token task (autoregressive models): From a standard causal forward pass, the representation is the hidden state (at a specified layer) at the last token *preceding* the predicted word, and the class label is the lemma of the actual next word, restricted to the same open-class content-word construction as above.

Audio - VGGSound. For audio we use VGGSound [Chen et al., 2020], a large-scale dataset of ~ 10 -second clips sourced from web video. We evaluate on the train split: 310 single-label classes and 183,630 clips (~ 592 for each class), decoded to 16 kHz mono.

Table 6: Model roster (36 backbones). N is the feature dimension of the extracted representation; P is the number of classes in the task.

Family	Model	N	Data / task	P	Objective
<i>Vision — ImageNet-1K</i>					
Vision	DINOv3 ViT-S/16	384	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ViT-S+/16	384	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ViT-B/16	768	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ViT-L/16	1024	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ViT-H+/16	1280	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ViT-7B/16	4096	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ConvNeXt-T	768	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ConvNeXt-S	768	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ConvNeXt-B	1024	ImageNet-1K	1000	Self-supervised
Vision	DINOv3 ConvNeXt-L	1536	ImageNet-1K	1000	Self-supervised
Vision	DINOv2 ViT-g/14	1536	ImageNet-1K	1000	Self-supervised
Vision	CLIP ViT-L/14	768	ImageNet-1K	1000	Contrastive (image-text)
Vision	ViT-B/16 (IN21K)	768	ImageNet-1K	1000	Supervised (IN21K)
Vision	ConvNeXt-B (IN1K)	1024	ImageNet-1K	1000	Supervised (IN1K)
Vision	ResNet-50	2048	ImageNet-1K	1000	Supervised (IN1K)
<i>Audio — VGGSound</i>					
Audio	SSLAM	768	VGGSound	310	Self-supervised + AudioSet-2M FT
Audio	CED-base	768	VGGSound	310	Supervised distillation (AudioSet)
Audio	AST	768	VGGSound	310	Supervised (AudioSet)
<i>Masked LM — TinyStories masked-word</i>					
Masked LM	RoBERTa-large	1024	TinyStories (masked)	320	Masked language modeling
Masked LM	BERT-large	1024	TinyStories (masked)	320	Masked language modeling
Masked LM	ModernBERT-large	1024	TinyStories (masked)	320	Masked language modeling
Masked LM	NomicBERT	768	TinyStories (masked)	320	Masked language modeling
<i>Autoregressive LM — TinyStories next-token</i>					
Autoregressive LM	Qwen2.5-0.5B	896	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen2.5-1.5B	1536	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen2.5-3B	2048	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen2.5-7B	3584	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen2.5-14B	5120	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen2.5-32B	5120	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-0.6B	1024	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-1.7B	2048	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-4B	2560	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-8B	4096	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-14B	5120	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Qwen3-32B	5120	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	Llama-3.1-8B	4096	TinyStories (next-token)	320	Causal language modeling
Autoregressive LM	OLMo-2-13B	5120	TinyStories (next-token)	320	Causal language modeling

B.2 Models

The canonical roster comprises 36 pretrained backbones across the four modality groups (Table 6). Vision contains the self-supervised DINOv3 family [Siméoni et al., 2025]: DINOv3 ViT-S/16 through ViT-7B/16 and DINOv3 ConvNeXt-T through ConvNeXt-L, together with the self-supervised DINOv2 ViT-g/14 [Oquab et al., 2023], contrastive CLIP ViT-L/14 [Radford et al., 2021], the supervised ImageNet-21K ViT-B/16 [Dosovitskiy et al., 2020], the supervised ImageNet-1K ConvNeXt-B [Liu et al., 2022], and the supervised ResNet-50 [He et al., 2016].

The language backbones are evaluated on the TinyStories dataset [Eldan and Li, 2023]. The masked-word encoders are RoBERTa-large [Liu et al., 2019b], BERT-large [Devlin et al., 2019], ModernBERT-large [Warner et al., 2025], and NomicBERT [Nussbaum et al., 2024]. The next-token autoregressive models are the Qwen2.5-Instruct family (0.5B–32B) [Qwen et al., 2025], the Qwen3 family (0.6B–32B) [Yang et al., 2025], Llama-3.1-8B [Grattafiori et al., 2024], and OLMo-2-13B [OLMo et al., 2024].

The audio backbones are SSLAM [Alex et al., 2025], an EAT-style self-supervised model fine-tuned on AudioSet-2M; CED-base [Dinkel et al., 2024], an AudioSet-supervised distilled transformer; and AST [Gong et al., 2021], an AudioSet-supervised spectrogram transformer.

B.3 Representation extraction

For vision transformers, we take the class token of the final layer as the representation. For convolutional architectures (ConvNeXt, ResNet), we take the global-average-pooled feature vector of the final stage. For audio backbones, we take the model’s pre-classifier pooled embedding: SSLAM its fc-norm of the class token, CED-base the mean over its final-layer hidden states, AST its standard pre-classifier pooled outputs. For masked language models the representation is the final hidden state at the masked position — the input to the masked-LM head. For autoregressive language models the representation is the hidden state (in a specific layer, see Sec. G) of the last token preceding the predicted word. In every case, the feature dimensions N reported in Table 6 are those of the stored arrays.

Before any analysis, we z-score each representation dimension-wise, using the empirical mean and standard deviation of the data.

B.4 Resources

Feature extraction ran on GPU (A100 and H100); the numerical integration of the theory and the fitting of the rescaling factor λ ran on CPU. The largest extractions were Qwen-14B/32B and OLMo on language, and DINOv3 ViT-7B on vision.

C Theoretical Proofs

This appendix contains all material deferred from Sec. 3 and Sec. 5. Appendix C.1 works out the isotropic uncorrelated model and derives the corollaries that show how the model fails qualitatively on real data. Appendix C.2 gives the full proofs of the centroid-aligned theory used in the main text.

C.1 Isotropic model

We analyze the maximally unstructured baseline, in which class centroids and within-class fluctuations are both isotropic Gaussian. It retains only the ambient dimension N , the number of classes P , and the noise-to-signal scale R , discarding every correlation between centroids and fluctuations. We derive its exact large- N accuracy, show that accuracy is self-averaging, and extract corollaries delineating its regimes. In Section 4 these predictions are shown to match real representations only after the residual-centroid correlations are removed.

Assumption C.1.1 (*Isotropic uncorrelated model*). For each class $\mu = 1, \dots, P$ the centroid is drawn $\mathbf{c}_\mu \sim_{\text{iid}} \mathcal{N}(\mathbf{0}, I_N)$ and held fixed (quenched). Conditioned on $\mathbf{c}_k$, an example of class k is $\mathbf{x} = \mathbf{c}_k + \delta\mathbf{x}$ with $\delta\mathbf{x} \sim \mathcal{N}(\mathbf{0}, R^2 I_N)$, sampled independently across examples.

Assumption C.1.2 (*High-dimensional regime*). $N \rightarrow \infty$, the number of classes grows at most polynomially, $P = N^{\mathcal{O}(1)}$, and the radius $R = R(N)$ may scale with N .

We refer to Assumptions C.1.1–C.1.2 jointly as the *isotropic model*. The classifier is the unit-centroid (prototype) rule $\hat{y}(\mathbf{x}) = \arg \max_\mu \hat{\mathbf{c}}_\mu \cdot \mathbf{x}$, with $\hat{\mathbf{c}}_\mu := \mathbf{c}_\mu / \|\mathbf{c}_\mu\|$, and its logits are $t_\mu := \hat{\mathbf{c}}_\mu \cdot \mathbf{x}$.

Lemma C.1.3 (*Overlaps*). Let $\mathbf{u}, \mathbf{v} \sim_{\text{iid}} \mathcal{N}(\mathbf{0}, I_N)$. Then, as $N \rightarrow \infty$, $\|\mathbf{u}\|/\sqrt{N} \xrightarrow{P} 1$ and $\sqrt{N}(\hat{\mathbf{u}} \cdot \hat{\mathbf{v}}) \xrightarrow{d} \mathcal{N}(0, 1)$. Consequently, the dot product of the target centroid with the true and rival directions obeys

$$\hat{\mathbf{c}}_\mu \cdot \mathbf{c}_k = \varepsilon_\mu + \delta_{\mu k}(\sqrt{N} - \varepsilon_\mu), \quad \varepsilon_\mu \sim_{\text{iid}} \mathcal{N}(0, 1). \quad (9)$$

Proof. $\mathbf{u} \cdot \mathbf{v} = \sum_i u_i v_i$ is a sum of N i.i.d. centred unit-variance variables, so $\mathbf{u} \cdot \mathbf{v} / \sqrt{N} \xrightarrow{d} \mathcal{N}(0, 1)$ by the CLT, while $\|\mathbf{u}\|/\sqrt{N} \xrightarrow{P} 1$ by the law of large numbers. $\square$

Theorem C.1.4 (*Mean accuracy*). *Under the isotropic model,*

$$\overline{\mathbb{E}_{\delta\mathbf{x}}[\text{Acc}]} = \mathbb{E}_{z \sim \mathcal{N}(0, 1)} \left[\Phi^{P-1} \left(\frac{\sqrt{N} + Rz}{\sqrt{1 + R^2}} \right) \right]. \quad (10)$$

where $\overline{\mathbb{E}_{\delta\mathbf{x}}[\text{Acc}]}$ denotes the average over the quenched disorder $\{\mathbf{c}_\mu\}$ after averaging over the thermal $\delta\mathbf{x}$.

Proof. Conditional on the example and centroids, the accuracy is the probability that the correct logit is the largest,

$$\text{Acc} \mid \delta\mathbf{x}, \{\mathbf{c}_\nu\} = \int_{-\infty}^{\infty} dt_k P(t_k) \prod_{\mu \neq k} \int_{-\infty}^{t_k} dt_\mu P(t_\mu), \quad P(t_\mu) = \delta(t_\mu - \hat{\mathbf{c}}_\mu \cdot \mathbf{x}). \quad (11)$$

Writing each δ through its Fourier representation introduces conjugate variables $\hat{t}_\mu$:

$$\text{Acc} \mid \delta\mathbf{x}, \{\mathbf{c}_\nu\} = \int_{-\infty}^{\infty} \frac{dt_k d\hat{t}_k}{2\pi} \prod_{\mu \neq k} \int_{-\infty}^{t_k} \frac{dt_\mu d\hat{t}_\mu}{2\pi} \exp\left(i \sum_{\mu=1}^P \hat{t}_\mu (t_\mu - \hat{\mathbf{c}}_\mu \cdot (\mathbf{c}_k + \delta\mathbf{x}))\right). \quad (12)$$

Carrying out the Gaussian $\delta\mathbf{x}$ and $\hat{t}_\mu$ integrals,

$$\mathbb{E}_{\delta\mathbf{x}}[\text{Acc} \mid \{\mathbf{c}_\mu\}] = \int_{-\infty}^{\infty} \frac{dt_k}{\sqrt{2\pi R^2}} \prod_{\mu \neq k} \int_{-\infty}^{t_k} \frac{dt_\mu}{\sqrt{2\pi R^2}} \exp\left(-\frac{1}{2R^2} \sum_{\mu=1}^P (t_\mu - \hat{\mathbf{c}}_\mu \cdot \mathbf{c}_k)^2\right). \quad (13)$$

Lemma C.1.3 sets $\hat{\mathbf{c}}_\mu \cdot \mathbf{c}_k = \varepsilon_\mu + \delta_{\mu k}(\sqrt{N} - \varepsilon_\mu)$. The target mean concentrates at $\sqrt{N}$, while integrating each rival overlap $\varepsilon_\mu \sim \mathcal{N}(0, 1)$ inflates the rival variance $R^2 \rightarrow 1 + R^2$:

$$\overline{\text{Acc}} = \int_{-\infty}^{\infty} \frac{dt_k}{\sqrt{2\pi R^2}} \prod_{\mu \neq k} \int_{-\infty}^{t_k} \frac{dt_\mu}{\sqrt{2\pi(1+R^2)}} \exp\left(-\frac{(t_k - \sqrt{N})^2}{2R^2} - \frac{1}{2} \sum_{\mu \neq k} \frac{t_\mu^2}{1+R^2}\right). \quad (14)$$

The rival integrals factorize into $\Phi(t_k/\sqrt{1+R^2})$; substituting $t_k = \sqrt{N} + Rz$ yields Eq. 10. $\square$

The average Eq. 10 is taken over both variability $\delta\mathbf{x}$ and centroids; the next result shows that the accuracy is self-averaging, such that taking the annealed average over the centroids disorder is valid.

Proposition C.1.5 (Self-averaging). Under the isotropic model, the accuracy is self-averaging. $\text{Var}(\text{Acc}) = \mathcal{O}(1/(PR^2)) \rightarrow 0$.

Proof. We bound the variance of the disorder-averaged accuracy. Introduce two independent replicas $\mathbf{x}, \mathbf{x}'$ of class k sharing the centroids. As in Eq.13, we get,

$$\mathbb{E}_{\delta\mathbf{x}}[\text{Acc}^2 \mid \{\mathbf{c}_\mu\}] = \int_{-\infty}^{\infty} \frac{dt_k dt'_k}{2\pi R^2} \prod_{\mu \neq k} \int_{-\infty}^{t_k} \frac{dt_\mu}{\sqrt{2\pi R^2}} \int_{-\infty}^{t'_k} \frac{dt'_\mu}{\sqrt{2\pi R^2}} \exp\left(-\frac{1}{2R^2} \sum_{\mu} [(t_\mu - \hat{\mathbf{c}}_\mu \cdot \mathbf{c}_k)^2 + (t'_\mu - \hat{\mathbf{c}}_\mu \cdot \mathbf{c}_k)^2]\right). \quad (15)$$

Averaging over the centroids couples the replicas through the shared mean. The integral becomes bivariate normal with variance $1 + R^2$ and covariance 1,

$$\overline{\text{Acc}^2} = \int \frac{dt_k dt'_k}{2\pi R^2} \prod_{\mu \neq k} \frac{1}{\mathcal{Z}} \int_{-\infty}^{t_k + \sqrt{N}} dt_\mu \int_{-\infty}^{t'_k + \sqrt{N}} dt'_\mu \exp\left(-\frac{t_k^2 + t'^2_k}{2R^2} - \frac{(1+R^2)(t_\mu^2 + t'^2_\mu) - 2t_\mu t'_\mu}{2R^2(R^2 + 2)}\right), \quad (16)$$

with the normalization $\mathcal{Z} = 2\pi R\sqrt{R^2 + 2}$.

The regime $R = \mathcal{O}(1)$: The target shift $\sqrt{N} \rightarrow \infty$ pushes the rival limits to $+\infty$, the rival integrals yield 1 up to an exponentially small correction whenever $\ln P = o(N)$, and $\overline{\text{Acc}^2} = \overline{\text{Acc}}^2 = 1$, so $\text{Var} \rightarrow 0$.

The regime $R \sim \sqrt{N}$: Set $r = R/\sqrt{N}$ and rescale. The inter-replica coupling is weak, of order $1/(Nr^2)$, and we can expand in the coupling. Solving the uncoupled Gaussian integrals yields:

$$\begin{aligned} \overline{\text{Acc}^2} &= \mathbb{E}_{z, z'} \left[\left(\Phi\left(z + \frac{1}{r}\right) \Phi\left(z' + \frac{1}{r}\right) + \frac{1}{Nr^2} \phi\left(z + \frac{1}{r}\right) \phi\left(z' + \frac{1}{r}\right) + \mathcal{O}(N^{-2}) \right)^{P-1} \right] \\ &= \overline{\text{Acc}}^2 + \frac{P-1}{Nr^2} \left(\mathbb{E}_z \left[\Phi^{P-2}\left(z + \frac{1}{r}\right) \phi\left(z + \frac{1}{r}\right) \right] \right)^2 + \mathcal{O}\left(\frac{\ln P}{N^2 P^2}\right). \end{aligned} \quad (17)$$

where $\phi(x)$ is a standard normal PDF, and $\Phi(x)$ its CDF. The expectation can be performed using the Laplace method, yielding the classic extreme-value extremum $z^* + 1/r = \sqrt{2 \ln P}$, $\phi(z^* + 1/r) = \sqrt{2 \ln P}/P$, and the width of the Laplace region $\Delta z = 1/(2\sqrt{\ln P})$. We note that the higher binomial terms are suppressed not by the $1/N$ prefactors but by the Gaussian PDF densities ϕ , each one of the order $\sqrt{2 \ln P}/P$, which cancel the binomial growth.

$$\text{Var}(\text{Acc}) = \frac{P-1}{Nr^2} \left(\mathbb{E}_z \left[\Phi^{P-2} \left(z + \frac{1}{r} \right) \phi \left(z + \frac{1}{r} \right) \right] \right)^2 = \mathcal{O} \left(\frac{1}{PNr^2} \right) = \mathcal{O} \left(\frac{1}{PR^2} \right) \rightarrow 0 \quad (18)$$

In both regimes $\text{Var}(\text{Acc}) \rightarrow 0$. $\square$

The mean accuracy Eq.10 has four immediate consequences, delineating the regimes of the isotropic model.

Corollary C.1.6 (*Separable regime*). For $R = \mathcal{O}(1)$ the classes are well separated and classification is trivial: $\text{Acc} \rightarrow 1$.

Proof. The scale of the maximum of P i.i.d. Gaussians with variance $\sigma^2 = 1 + R^2$ converges to $\sigma\sqrt{2 \ln P}$ for large P . The signal is $\mathbb{E}[t_k] = \sqrt{N}$. From Assumption C.1.2, $\ln P = \mathcal{O}(\ln N)$ thus for any $R = \mathcal{O}(1)$, $\sigma\sqrt{2 \ln P} \ll \sqrt{N}$ which yields $\text{Acc} \rightarrow 1$. $\square$

Corollary C.1.7 (*Saturation regime*). In the infinite-radius limit $R/\sqrt{N} \rightarrow \infty$ the isotropic residual dominates the input, which retains no information about the true class: all P classes are equally likely and accuracy collapses to chance, $\text{Acc} \rightarrow 1/P$.

Proof. Dividing through by R , the argument of Φ in Eq.10 tends to z pointwise, so bounded convergence gives $\text{Acc} \rightarrow \mathbb{E}_z[\Phi^{P-1}(z)] = \mathbb{P}(\xi_k > \max_{\mu \neq k} \xi_\mu) = 1/P$ for $\xi_1, \dots, \xi_P \sim_{\text{iid}} \mathcal{N}(0, 1)$, by permutation symmetry. $\square$

Corollary C.1.8 (*Critical radius*). Non-perfect accuracy requires large radii:

$$\text{Acc} < 1 \Rightarrow R = \Omega(\sqrt{N/\ln P}).$$

Proof. Contrapositive. If $R = o(\sqrt{N/\ln P})$ then $1 + R^2 = o(N/\ln P)$, so the argument of Φ in Eq.10 is $(\sqrt{N} + Rz)/\sqrt{1 + R^2} = \omega(\sqrt{\ln P}) \gg \sqrt{2 \ln P}$, and Corollary C.1.6 forces $\text{Acc} \rightarrow 1$. $\square$

Corollary C.1.9 (*Per-class fluctuations*). Per-class accuracy fluctuations are of order $1/R\sqrt{P}$.

Proof. Calculation of the variance in Proposition C.1.5 yields $\text{Var}(\text{Acc}) = \mathcal{O}(1/PR^2)$. Different classes are equivalent to a resampling of the centroids, so per-class fluctuations are of the same order. $\square$

C.2 Centroid-aligned model

Definition C.2.1 (*Centroid geometry*). Fix the true class k . For $\mu, \nu \in \{1, \dots, P\}$ let $G_{\mu\nu} := \hat{\mathbf{c}}_\mu \cdot \hat{\mathbf{c}}_\nu$ be the Gram matrix of the unit centroids, and let $g_\mu := \hat{\mathbf{c}}_k \cdot \hat{\mathbf{c}}_\mu$ be the centroid overlap of class μ with the true class, so that $g_k = 1$. The rival support is $\mathcal{R}_K$ and $\mathcal{R}_K^+ := \{k\} \cup \mathcal{R}_K$, as defined in Sec. 5.

Assumption C.2.2 (*Deterministic centroids and Gaussian projections*). We assume that the centroids $\{\mathbf{c}_\mu\}_{\mu=1}^P$ are deterministic and held fixed, and that the restricted Gram matrix of $\mathcal{R}_K^+$ is invertible. The source of stochasticity in the system is the sampling of Gaussian projection coefficients $\{s_\mu\}_{\mu=1}^P$ and $s_\perp$.

Definition C.2.3 (*Centroid-aligned generative model*). An example of class k is generated by

$$\mathbf{x}(\mathbf{s}) = \|\mathbf{c}_k\| \left[\hat{\mathbf{c}}_k + R \left(\sum_{\mu, \nu \in \mathcal{R}_K^+} \sigma_\mu s_\mu G_{\mu\nu}^{-1} \hat{\mathbf{c}}_\nu + \sigma_\perp s_\perp \hat{\mathbf{e}}_\perp \right) \right], \quad (19)$$

where $\hat{\mathbf{e}}_\perp$ is orthogonal to $\text{span}(\hat{\mathbf{c}}_\mu)_{\mu \in \mathcal{R}_K^+}$, R is the empirical class radius, and σ_μ is the projected standard deviation along $\hat{\mathbf{c}}_\mu$.

Definition C.2.4 (*Normalized logits*). The prototype logits of (Def 2.1) are defined as $t_\mu = \hat{\mathbf{c}}_\mu \cdot \mathbf{x}$. Throughout this section it is convenient to divide all of them by the common factor $\|\mathbf{c}_k\|$, and we set

$$t_\mu := \frac{\mathbf{x} \cdot \hat{\mathbf{c}}_\mu}{\|\mathbf{c}_k\|}, \quad \mu \in \mathcal{R}_K^+. \quad (20)$$

Since $\|\mathbf{c}_k\|$ is a common factor independent of μ , the argmax properties of the logits do not change. The classifier and its accuracy are therefore those of Definition 2.1, and no generality is lost. We restrict the logits only to the sparse set of rivals $\mathcal{R}_K$, and neglect weak competitors.

Lemma C.2.5 (*Logit representation*). Under Assumption C.2.2 and Definitions C.2.3–C.2.4, $t_\mu = g_\mu + R \sigma_\mu s_\mu$ for every $\mu \in \mathcal{R}_K^+$.

Proof. Taking the inner product of Eq. 19 with $\hat{\mathbf{c}}_\mu$ and dividing by $\|\mathbf{c}_k\|$ gives $t_\mu = g_\mu + R(\sum_{\alpha, \nu \in \mathcal{R}_K^+} \sigma_\alpha s_\alpha G_{\alpha\nu}^{-1} G_{\nu\mu} + \sigma_\perp s_\perp \hat{\mathbf{e}}_\perp \cdot \hat{\mathbf{c}}_\mu)$. The last term vanishes since $\hat{\mathbf{e}}_\perp \perp \hat{\mathbf{c}}_\mu$ for $\mu \in \mathcal{R}_K^+$, and $\sum_\nu G_{\alpha\nu}^{-1} G_{\nu\mu} = \delta_{\alpha\mu}$ by Assumption C.2.2. $\square$

Definition C.2.6 (*Projection covariance*). The coefficients are centered and standardized by definition, $\mathbb{E}_k[s_\mu] = 0$ and $\mathbb{E}_k[s_\mu^2] = 1$. We write

$$\mathbb{E}_k[s_\mu s_\nu] = \frac{\Sigma_{\mu\nu}}{R^2 \sigma_\mu \sigma_\nu}, \quad \mu, \nu \in \mathcal{R}_K^+. \quad (21)$$

By Lemma C.2.5 this makes $\Sigma_{\mu\nu} = \text{Cov}_k[t_\mu, t_\nu]$ the covariance matrix of the logit vector itself. No structure is imposed on Σ beyond that of a well-conditioned covariance matrix.

Proposition C.2.7 (*Accuracy as an orthant probability*). Under Lemma C.2.5, Definition C.2.6 and Assumption C.2.2,

$$\text{Acc}^{\text{th}}(R) = \frac{1}{\sqrt{\det \Sigma}} \int_{-\infty}^{\infty} \frac{dt_k}{\sqrt{2\pi}} \prod_{\mu \in \mathcal{R}_K} \int_{-\infty}^{t_k} \frac{dt_\mu}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} \sum_{\mu, \nu \in \mathcal{R}_K^+} (t_\mu - g_\mu) \Sigma_{\mu\nu}^{-1} (t_\nu - g_\nu)\right). \quad (22)$$

Proof. By Lemma C.2.5, Definition C.2.6 and Assumption C.2.2 the vector $(t_\mu)_{\mu \in \mathcal{R}_K^+}$ is Gaussian with mean (g_μ) and covariance Σ . An example is classified correctly iff $t_\mu < t_k$ for all $\mu \in \mathcal{R}_K$. $\square$

Eq. 22 is a multivariate orthant probability, which in general can be evaluated only numerically (for example by Monte-Carlo simulation). A specific structure admits an analytical solution.

Lemma C.2.8 (*Gaussian conditioning*). Under Assumption C.2.2, set $\rho_\mu := \mathbb{E}_k[s_\mu s_k]$ for $\mu \in \mathcal{R}_K$. Then for all $\mu, \nu \in \mathcal{R}_K$,

$$\mathbb{E}_k[t_\mu | t_k] = g_\mu + \rho_\mu \sigma_\mu \frac{t_k - 1}{\sigma_k}, \quad \text{Cov}[t_\mu, t_\nu | t_k] = R^2 \sigma_\mu \sigma_\nu (\mathbb{E}_k[s_\mu s_\nu] - \rho_\mu \rho_\nu), \quad (23)$$

and in particular the conditional covariance does not depend on t_k .

Proof. By Definition C.2.6, $\text{Cov}(t_\mu, t_\nu) = \Sigma_{\mu\nu}$, $\Sigma_{\mu k} = R^2 \sigma_\mu \sigma_k \rho_\mu$ and $\Sigma_{kk} = R^2 \sigma_k^2$, while $\mathbb{E}_k[t_\mu] = g_\mu$ and $\mathbb{E}_k[t_k] = 1$. For a jointly Gaussian vector, $\mathbb{E}[t_\mu | t_k] = \mathbb{E}[t_\mu] + \Sigma_{\mu k} \Sigma_{kk}^{-1} (t_k - \mathbb{E}[t_k])$ and $\text{Cov}[t_\mu, t_\nu | t_k] = \Sigma_{\mu\nu} - \Sigma_{\mu k} \Sigma_{\nu k} \Sigma_{kk}^{-1}$. Substituting gives Eq. 23. $\square$

Assumption C.2.9 (*Conditional independence*). Conditionally on t_k the rival logits are uncorrelated: $\text{Cov}[t_\mu, t_\nu | t_k] = 0$ for all $\mu \neq \nu$ in $\mathcal{R}_K$. Equivalently, by Lemma C.2.8,

$$\mathbb{E}_k[s_\mu s_\nu] = \delta_{\mu\nu} (1 - \rho_\mu^2) + \rho_\mu \rho_\nu, \quad \text{so that} \quad \text{Cov}[t_\mu, t_\nu | t_k] = \delta_{\mu\nu} R^2 \sigma_\mu^2 (1 - \rho_\mu^2). \quad (24)$$

Theorem C.2.10 (*Conditional factorization*). Under Assumption C.2.9,

$$\text{Acc}^{\text{th}}(R) = \int_{-\infty}^{\infty} \frac{dt_k}{R\sigma_k\sqrt{2\pi}} \exp\left(-\frac{(t_k-1)^2}{2R^2\sigma_k^2}\right) \prod_{\mu \in \mathcal{R}_K} \Phi\left(\frac{t_k(1-\rho_\mu\sigma_\mu/\sigma_k) + \rho_\mu\sigma_\mu/\sigma_k - g_\mu}{R\sigma_\mu\sqrt{1-\rho_\mu^2}}\right). \quad (25)$$

Proof. Condition on t_k . By Lemma C.2.8 and Assumption C.2.9 the variables $(t_\mu)_{\mu \in \mathcal{R}_K}$ are then independent Gaussians with mean $g_\mu + \rho_\mu\sigma_\mu(t_k-1)/\sigma_k$ and variance $R^2\sigma_\mu^2(1-\rho_\mu^2)$, so the orthant probability factorizes into $\prod_{\mu} \Pr[t_\mu < t_k \mid t_k]$, which yields Eq. 25. $\square$

Corollary C.2.11 (*Standardized form*). Substituting $t_k = R\sigma_k z + 1$ in Eq. 25 yields

$$\text{Acc}_k^{\text{th}}(R) = \mathbb{E}_{z \sim \mathcal{N}(0,1)} \left[\prod_{\mu \in \mathcal{R}_K} \Phi\left(\frac{\sigma_k - \rho_\mu\sigma_\mu}{\sigma_\mu\sqrt{1-\rho_\mu^2}} z + \frac{1-g_\mu}{R\sigma_\mu\sqrt{1-\rho_\mu^2}}\right) \right], \quad (26)$$

which is Theorem 5.5 of the main text.

Proposition C.2.12 (*Unconditional moments*). Under Assumption C.2.9, for $\mu, \nu \in \mathcal{R}_K$,

$$\mathbb{E}_k[t_\mu] = g_\mu, \quad \Sigma_{\mu\nu} = \text{Cov}_k[t_\mu, t_\nu] = R^2 (\delta_{\mu\nu}\sigma_\mu^2(1-\rho_\mu^2) + \sigma_\mu\sigma_\nu\rho_\mu\rho_\nu), \quad (27)$$

which is a rank-one correction to a diagonal covariance. This is the covariance structure of $\mathbf{s}$ stated in Assumption 5.4 of the main text.

Proof. From Lemma C.2.5, $\mathbb{E}_k[t_\mu] = g_\mu$ and $\Sigma_{\mu\nu} = R^2\sigma_\mu\sigma_\nu\mathbb{E}_k[s_\mu s_\nu]$. Inserting the covariance structure from Eq. 24 yields the result. $\square$

Remark C.2.13 (*Quenched versus deterministic centroids*). Assumption C.2.2 treats the centroids as deterministic. Drawing them instead from a Gaussian ensemble whose second moments reproduce the overlap structure G does not change the result for large N : the induced fluctuations of the coefficients g_μ , σ_μ and ρ_μ entering Eq. 25 are subleading corrections to their $\mathcal{O}(1)$ values, and the accuracy can be shown to be self-averaging by the same argument as in Proposition C.1.5.

D Synthetic Data

We validate the two generative models of the main text on synthetic data, where the geometry is controlled exactly and the analytic predictions can be checked without any of the confounds of real representations. Figure 6 tests the isotropic uncorrelated model of Section 3, and Figure 7 tests the centroid-aligned model of Section 5. For each model we verify three claims: (i) the analytic accuracy matches the simulation; (ii) the accuracy is self-averaging, i.e. a single quenched centroid bank gives the same per-class accuracy as a fresh, independently drawn set of rivals for every example; and (iii) the fluctuation of the accuracy across realizations shrinks with the number of classes P and for the isotropic model also with the class radius R as predicted by the previous section.

E Rival Selection

For a given example, only a small number of the P classes are genuine competitors. This is the reason that the theory of Section 5 evaluates the accuracy over a small rival set $\mathcal{R}_K$. This appendix (i) compares four possible rival-selection rules on convergence to the full P -way classification (ii) shows that the theory’s error saturates by $K \approx 10$ -20, so $K = 20$ is a principled choice.

E.1 Rival-selection rules

Fix a class k . We define a ranking method for every candidate rival $\mu \neq k$:

Centroid correlations: $g_\mu = \hat{\mathbf{c}}_k \cdot \hat{\mathbf{c}}_\mu$, (descending).

Projection variance: $\sigma_\mu = \sqrt{\mathbb{E}_k[(\delta\mathbf{x} \cdot \hat{\mathbf{c}}_\mu)^2] / \mathbb{E}_k[\|\delta\mathbf{x}\|^2]}$, (descending).

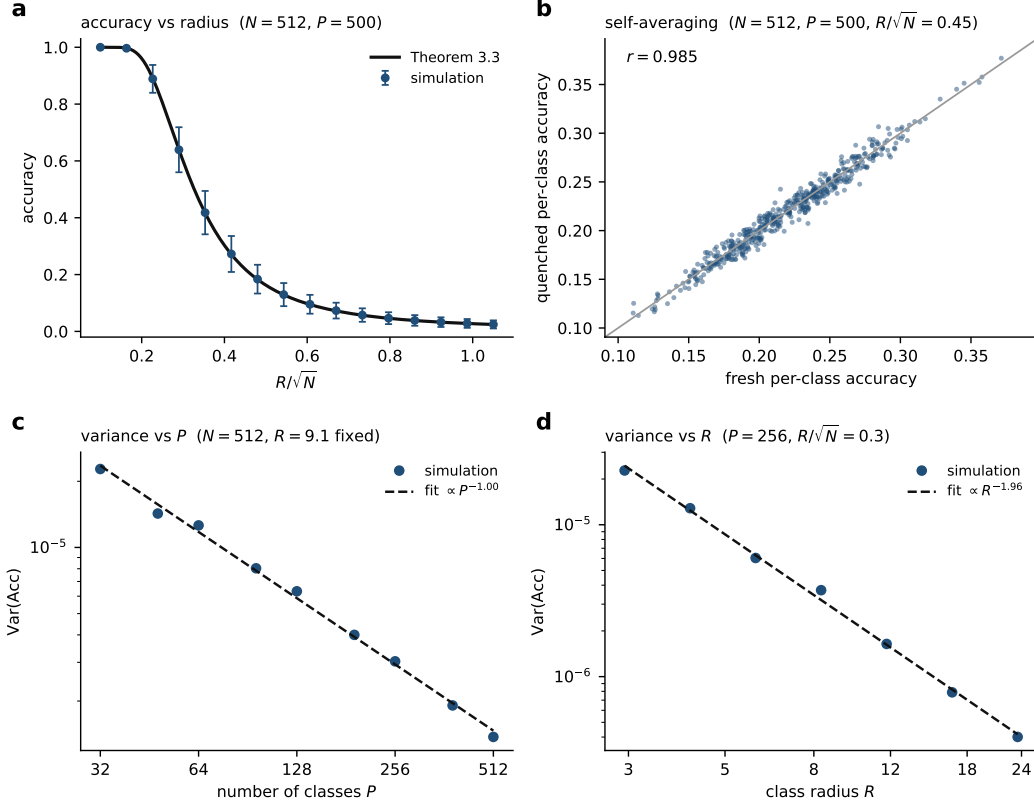


Figure 6: **Validation of the isotropic model’s theory on synthetic data.** (a) Class-averaged accuracy versus $R/\sqrt{N}$: simulation (markers, mean $\pm$ between-class std) against Theorem 3.3 (line), $N = 512$, $P = 500$. (b) Self-averaging: quenched per-class accuracy (one shared centroid bank) versus the fresh-rival accuracy (rivals redrawn per example), the line is $y = x$. $r = 0.985$, confirming self-averaging - the thermal average is equivalent to the quenched average. (c) Across-realization variance of the accuracy versus P at fixed radius $R = 9.05$, fitted power law $\propto P^{-1.00}$, predicted by Proposition C.1.5. (d) Variance versus radius R (swept via the dimension at fixed $R/\sqrt{N}$), fitted power law $\propto R^{-1.96}$, predicted by Proposition C.1.5.

Theory margin: $(1 - g_\mu) / (R\sigma_\mu \sqrt{1 - \rho_\mu^2})$, (ascending). The z-independent argument in Eq.7.

Gaussian margin: $(1 - g_\mu) / \sqrt{\sigma_k^2 + \sigma_\mu^2 - 2\sigma_k\sigma_\mu\rho_\mu}$, (ascending). The exact z-expectation over Eq.7 for one rival.

E.2 Convergence of the K -way classification accuracy

Restricting the classification to a smaller rival set can only improve accuracy, so the K -way accuracy exceeds the full P -way accuracy and decreases monotonically to it as K grows. Table 7 reports, for each selector, the median number of rivals a backbone needs before its K -way accuracy is within a fixed band of the full accuracy, aggregated over all 36 backbones. The margin and variance selectors predict the P -way classification at the smallest K ; the centroid-correlation and Gaussian-margin selectors need substantially more rivals at the tight 1% band, confirming that ranking by projected variance (or by the conditional margin) is the most efficient way to choose the true competitors. The conclusion is consistent for all modalities. Since the projection variance and the conditional margin perform comparably (within five rivals at every band), we adopt the projection variance in the main text, as it is a single measured quantity and is therefore cheaper to compute and easier to interpret.

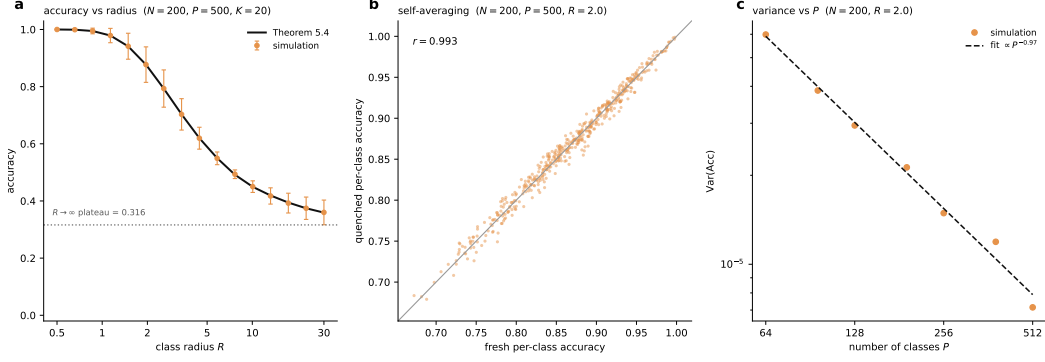


Figure 7: **Synthetic validation of the centroid-aligned model.** Examples are generated according to the centroid-aligned variability model (Sec. 5 and Sec. C.2) with $K = 20$ rivals, $N = 200$, $P = 500$. The geometry is calibrated to DINOv3 ViT-7B: own-class projected std $\sigma_k = 0.24$, rival stds $\sigma_\mu = 0.12\mu^{-0.4}$. **(a)** Class-averaged accuracy versus class radius R : simulation (markers) against Theorem 5.5 (line). The dotted line is the $R \rightarrow \infty$ plateau 0.316, which is set by σ_k/σ_μ alone. **(b)** Quenched centroid versus fresh per-example sampling of the centroids, $y = x$ line, $r = 0.993$, validating self-averaging. **(c)** Across-realization variance versus P at fixed radius, fitted $\propto P^{-0.97}$, similar to the isotropic model (Proposition C.1.5).

E.3 Optimal K for theory predictions

Increasing K enlarges the rival set the theory integrates over. Table 9 tracks the root-mean-square error between the predicted and empirical per-class accuracies as a function of K . The error drops steeply up to $K \approx 10$ and is essentially flat thereafter: at $K = 20$ the mean error is within 0.003 of its $K \rightarrow 300$ floor for 31 of the 36 backbones, and the median number of rivals needed to come within 0.001 of the floor is 20, the value we chose for K in the theory throughout.

Table 7: **Rivals needed to reach the full accuracy.** Median number of rivals K (over 36 backbones) at which the restricted K -way accuracy comes within 5/3/2/1% of the full P -way accuracy, per selector. Smaller is better.

Selector	within 5%	within 3%	within 2%	within 1%
Centroid correlations (g_μ)	14	24	36	87
Projection variance (σ_μ)	14	20	30	54
Theory margin	11	18	25	49
Gaussian margin	10	19	28	73

Table 8: **Rivals needed within 1%, by modality.** Median K to reach within 1% of the full accuracy, per selector and modality (n backbones).

Modality	n	Centroids corr.	Proj. variance	Theory margin	Gaussian margin
Vision	15	76	47	45	56
Autoregressive LM	14	102	89	82	108
Masked LM	4	48	42	35	43
Audio	3	60	39	37	49

Table 9: **The theory’s error saturates by $K \approx 10$ –20.** Mean per-class accuracy RMSE of the theory at $K = 10$ and $K = 20$, its large- K floor, and the number of backbones already within 0.003 of that floor (large K) at $K = 20$, by modality.

Modality	n	RMSE $K=10$	RMSE $K=20$	floor	within 0.003 by $K=20$
Vision	15	0.049	0.047	0.045	13/15
Audio	3	0.081	0.076	0.064	0/3
Masked LM	4	0.064	0.062	0.061	4/4
Autoregressive LM	14	0.080	0.075	0.075	14/14
All	36	0.065	0.062	0.060	31/36

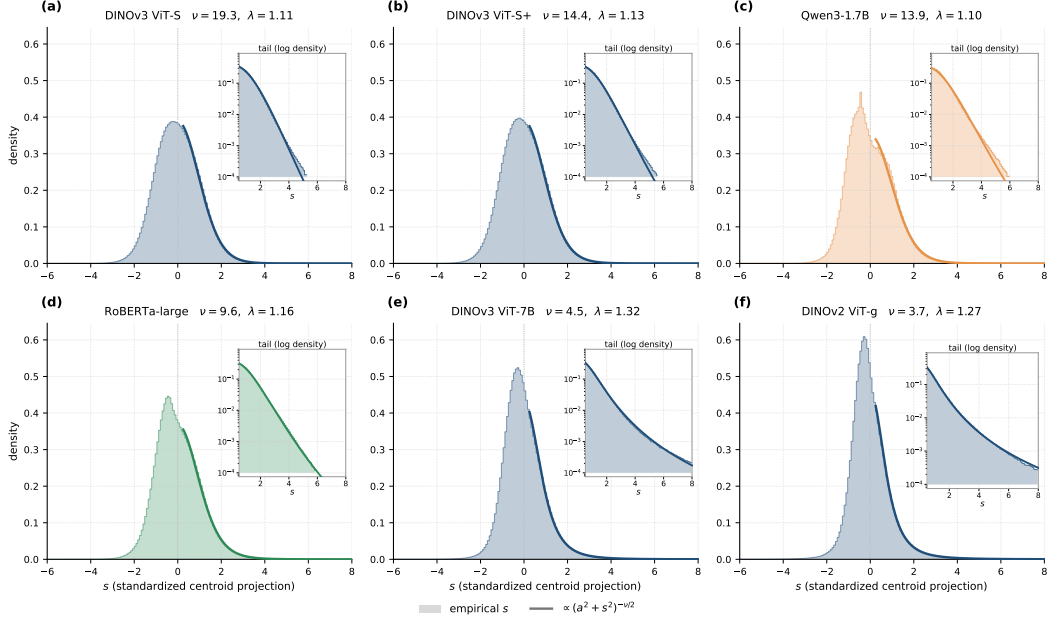


Figure 8: **Standardized centroid projections have a heavy right tail that varies across models.** Density of the standardized projection s (filled) with its fitted half-Student-t right tail $p(s) \propto (a^2 + s^2)^{-\nu/2}$ (solid), $K = 20$. Panels run light $\rightarrow$ heavy tail (left to right, top to bottom), with the fitted exponent ν and model in each title. Color encodes modality (vision blue, autoregressive LM orange, masked LM green). As ν decreases, the peak sharpens, and the right tail fattens.

F Heavy tail statistics and renormalization

The theory of Sect. 5 treats the standardized centroid projections s_μ as jointly Gaussian. Empirically, the variables s_μ (see Table. 2 for the definition) have a pronounced right tail. Since a prediction is decided by the largest rival score $\max_\mu t_\mu$, this positive tail controls the error rate, and it is exactly where a Gaussian underestimates the extreme value. Fitting the pooled positive projections gives a power-law

$$p(s) \propto (a^2 + s^2)^{-\nu/2} \quad (28)$$

obtained by a maximum-likelihood fit of a Student-t distribution of parameters a, ν to the positive s values. Here ν is the tail (power-law) exponent, a smaller ν means a heavier tail, and $\nu \rightarrow \infty$ recovers a Gaussian. A heavier tail inflates the extreme rival projections that produce errors, so the Gaussian theory is over-optimistic; the single global rescaling $R \rightarrow \lambda R$ used in Section 5 (fit per model) compensates for this by enlarging the effective within-class radius, and the fitted λ tracks the tail heaviness $1/\nu$ (see Fig.4 and Table 11).

Figure 8 contrasts the density of s across six backbones spanning three modalities, ordered from the lightest to the heaviest tail, each with its fitted curve (28).

In Sec. 5.2 we claim that without the additional rescaling of the variability by λ , the theory is overoptimistic. This is demonstrated explicitly in Table 10. The theory with $\lambda = 1$ consistently predicts accuracy high by 6% on average across all backbones, compared to only 0.5% with λ .

Table 11 lists the tail exponent ν for all 36 backbones, grouped by modality and sorted from heaviest to lightest tail, alongside the rescaling λ . The exponent varies substantially, from $\nu \approx 3.7$ (DINOv2 ViT-g, the heaviest tail on the roster) to $\nu \approx 22$ (Qwen3-0.6B, nearly Gaussian); the vision backbones span the widest range, while the audio and language backbones have larger ν (lighter tail) and less variability between models. Across the roster the heaviness $1/\nu$ correlates strongly with the fitted rescaling λ (Pearson $r = 0.90$ over the 36 backbones): the heaviest-tail models require the largest correction (λ up to 1.34) and the near-Gaussian ones need almost none. The main text shows this λ -versus- $1/\nu$ relationship as a scatter (Fig. 4).

Table 10: **Without global radius rescaling, the theory is over-optimistic.** Per-class accuracy predicted by the centroid-aligned theory of Sec. 5.2, evaluated at $\lambda = 1$ (no rescaling of the variability) and at the fitted λ , compared with the empirical accuracy. The theoretical accuracy with $\lambda = 1$ is higher than the empirical one, while the theory with λ is within 1% of the empirical one, with significantly better RMSE.

Modality	n	$\langle \text{Acc} \rangle$	$\lambda = 1$		fitted λ		$\langle \lambda \rangle$
			$\langle \text{Acc}^{\text{th}} \rangle$	RMSE	$\langle \text{Acc}^{\text{th}} \rangle$	RMSE	
Vision	15	0.812	0.876	0.080	0.819	0.047	1.228
Audio	3	0.613	0.657	0.087	0.617	0.076	1.084
Masked LM	4	0.715	0.785	0.090	0.724	0.062	1.131
Autoregressive LM	14	0.522	0.583	0.096	0.524	0.075	1.111
All	36	0.672	0.734	0.088	0.677	0.062	1.160

Backbone	ν	λ
<i>Vision (ImageNet-1K)</i>		
DINOv2 ViT-g	3.7	1.272
ConvNeXt-B	4.0	1.254
ViT-B	4.1	1.344
DINOv3 CNX-L	4.4	1.305
DINOv3 ViT-7B	4.5	1.316
DINOv3 CNX-B	4.5	1.296
ResNet50	4.6	1.228
DINOv3 CNX-S	5.2	1.229
DINOv3 ViT-H ⁺	5.3	1.297
DINOv3 CNX-T	6.8	1.148
DINOv3 ViT-L	7.0	1.221
DINOv3 ViT-B	9.0	1.158
DINOv3 ViT-S ⁺	14.4	1.134
CLIP-L	16.4	1.108
DINOv3 ViT-S	19.3	1.114
$r(\lambda, 1/\nu)$	0.888	
<i>Audio (VGGSound)</i>		
SSLAM	7.9	1.106
CED	7.9	1.089
AST	9.5	1.057
$r(\lambda, 1/\nu)$	0.948	

Backbone	ν	λ
<i>Masked LM (TinyStories, masked-word)</i>		
RoBERTa-large	9.6	1.157
BERT-large	10.1	1.127
ModernBERT-large	11.3	1.124
NomicBERT	14.2	1.115
$r(\lambda, 1/\nu)$		0.805
<i>Autoregressive LM (TinyStories, next-token)</i>		
OLMo-2-13B	10.2	1.152
Qwen2.5-32B	10.9	1.117
Qwen2.5-7B	11.2	1.109
Qwen3-8B	11.3	1.128
Qwen2.5-1.5B	11.3	1.110
Qwen3-14B	11.3	1.124
Qwen3-32B	11.4	1.119
Qwen2.5-14B	11.5	1.112
Llama-3.1-8B	12.4	1.107
Qwen2.5-3B	12.6	1.112
Qwen3-4B	12.7	1.118
Qwen3-1.7B	13.9	1.103
Qwen2.5-0.5B	19.2	1.088
Qwen3-0.6B	21.9	1.059
$r(\lambda, 1/\nu)$		0.887
Overall	0.901	

Table 11: **Right-tail exponent ν across all backbones.** ν is the power-law exponent of the fitted right tail $p(s) \propto (a^2 + s^2)^{-\nu/2}$; a smaller ν means a heavier tail. λ is the global radius rescaling of the theory. Grouped by modality, sorted from heaviest to lightest tail. Pearson correlations between λ and $1/\nu$ are high in every family ($r = 0.89, 0.95, 0.81, 0.89$ in vision, audio, masked language models and autoregressive models, respectively) and globally over all models ($r = 0.90$).

Backbone	P	N	Acc	Acc th	r	λ	RMSE
<i>Vision (ImageNet-1K)</i>							
ConvNeXt-B	1000	1024	0.943	0.952	0.954	1.254	0.0256
ResNet50	1000	2048	0.861	0.871	0.936	1.228	0.0454
DINOv3 ViT-H ⁺	1000	1280	0.856	0.862	0.960	1.297	0.0390
DINOv3 ViT-7B	1000	4096	0.855	0.861	0.951	1.316	0.0427
ViT-B	1000	768	0.853	0.861	0.959	1.344	0.0378
DINOv3 ViT-L	1000	1024	0.845	0.852	0.964	1.221	0.0384
DINOv3 CNX-L	1000	1536	0.814	0.819	0.961	1.305	0.0440
DINOv3 CNX-B	1000	1024	0.804	0.808	0.958	1.296	0.0455
DINOv2 ViT-g	1000	1536	0.795	0.804	0.941	1.272	0.0631
DINOv3 ViT-B	1000	768	0.794	0.800	0.954	1.158	0.0460
CLIP-L	1000	768	0.792	0.804	0.936	1.108	0.0596
DINOv3 CNX-S	1000	768	0.786	0.793	0.956	1.229	0.0481
DINOv3 CNX-T	1000	768	0.740	0.749	0.937	1.148	0.0602
DINOv3 ViT-S ⁺	1000	384	0.738	0.742	0.951	1.134	0.0504
DINOv3 ViT-S	1000	384	0.706	0.710	0.944	1.114	0.0555
<i>mean</i>			0.812	0.819	0.951	1.228	0.0468
<i>Audio (VGGSound)</i>							
SSLAM	310	768	0.628	0.631	0.952	1.106	0.0720
CED	310	768	0.628	0.632	0.947	1.089	0.0739
AST	310	768	0.583	0.587	0.944	1.057	0.0819
<i>mean</i>			0.613	0.617	0.948	1.084	0.0759
<i>Masked LM (TinyStories, masked-word)</i>							
RoBERTa-large	320	1024	0.767	0.776	0.944	1.157	0.0542
ModernBERT-large	320	1024	0.734	0.743	0.942	1.124	0.0614
BERT-large	320	1024	0.710	0.720	0.942	1.127	0.0630
NomicBERT	320	768	0.649	0.658	0.945	1.115	0.0683
<i>mean</i>			0.715	0.724	0.943	1.131	0.0617
<i>Autoregressive LM (TinyStories, next-token)</i>							
Qwen2.5-32B	320	5120	0.567	0.572	0.945	1.117	0.0722
Qwen3-32B	320	5120	0.566	0.569	0.955	1.119	0.0665
OLMo-2-13B	320	5120	0.563	0.567	0.948	1.152	0.0701
Qwen3-14B	320	5120	0.556	0.558	0.954	1.124	0.0694
Qwen2.5-14B	320	5120	0.552	0.556	0.942	1.112	0.0758
Llama-3.1-8B	320	4096	0.550	0.553	0.943	1.107	0.0772
Qwen2.5-7B	320	3584	0.544	0.550	0.940	1.109	0.0756
Qwen3-8B	320	4096	0.540	0.542	0.955	1.128	0.0693
Qwen2.5-3B	320	2048	0.528	0.531	0.946	1.112	0.0785
Qwen3-4B	320	2560	0.519	0.520	0.953	1.118	0.0697
Qwen2.5-1.5B	320	1536	0.491	0.494	0.931	1.110	0.0822
Qwen3-1.7B	320	2048	0.474	0.474	0.949	1.103	0.0747
Qwen2.5-0.5B	320	896	0.452	0.450	0.944	1.088	0.0853
Qwen3-0.6B	320	1024	0.406	0.403	0.939	1.059	0.0839
<i>mean</i>			0.522	0.524	0.946	1.111	0.0750
Overall			0.672	0.677	0.948	1.160	0.0618

Table 12: **Centroid-aligned theory**: The theory was tested on all 36 models. For each backbone, we report: number of classes P , feature dimension N , empirical prototype accuracy, per-class Pearson r between predicted and empirical accuracy, the fitted global rescaling λ , and the per-class prediction error RMSE. The theory consistently predicts the accuracy well across all models ($r > 0.93$).

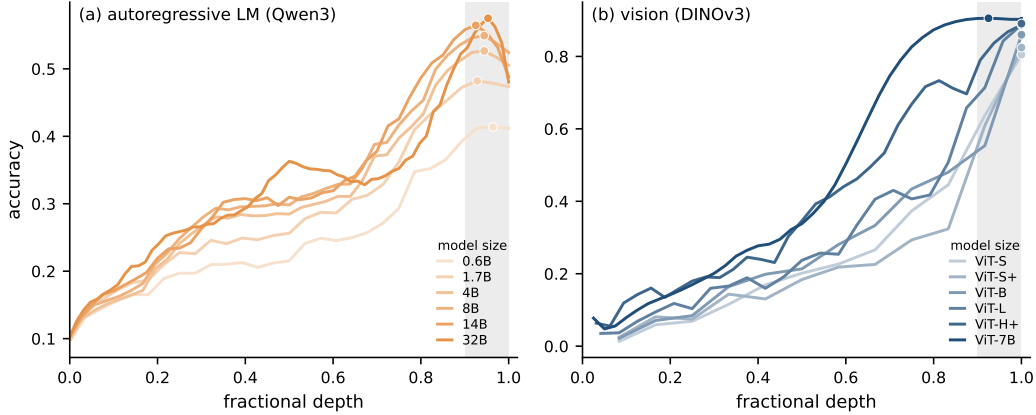


Figure 9: **The best prototype layer is a few blocks before the final one for language, but is the final block for vision.** Per-layer prototype classification accuracy versus fractional depth. Color runs from light (small) to dark (large) within each modality, and the peak performance is marked. **(a)** The six Qwen3 autoregressive language models; the optimal accuracy is a few blocks before the final layer, with a big drop at the final layer in most models. **(b)** the six DINOv3 vision transformers. The final layer is the best, with a plateau on DINOv3 7B.

G Depth analysis

The prototype classifier of Sec. 5 reads out a single hidden layer. Here we analyze how prototype classification accuracy and geometry change across depth, and identify the best layer for this task. For the autoregressive language models the best layer for content-word prediction is consistently a few blocks before the final one, and accuracy dips at the last block; for vision the final block is best (in DINOv3 7B, there is a plateau). We subsequently analyze the geometry of the largest language model, Qwen3-32B across depth, similar to the analysis in Sec. 6 for the Qwen3 family across scale. Similar to the scale effect, we see that the best layer has large true centroid correlation compared to rivals σ_k/σ_μ^* , decorrelates centroids (small g_μ^*), and has small correlations between projections ρ_μ^* . We note that, as observed before with scale, the class radius R does not predict the overall best accuracy.

Table 13: **Qwen3-32B across depth.** Class-averaged geometric measures for the seven last blocks of Qwen3-32B (64 blocks). * marks each class’s hardest competitor μ^* , measured by σ_μ , the projection standard deviation. $\langle R \rangle$ is the mean class radius; σ_k/σ_μ^* is the ratio between the standard deviation of the projection onto the true-class and hardest-rival directions. σ_μ^* is the hardest rival projected std, $g_\mu^* = \hat{c}_k \cdot \hat{c}_{\mu^*}$ the centroid overlap and ρ_μ^* the correlation between projections. ν is the right-tail exponent. Best measure highlighted in bold.

Layer	$\langle R \rangle$	σ_k/σ_μ^*	σ_μ^*	g_μ^*	ρ_μ^*	ν	Acc
L58 (0.91)	2.40	0.993	0.172	0.514	0.496	10.9	0.521
L59 (0.92)	2.32	0.996	0.175	0.511	0.502	11.0	0.542
L60 (0.94)	2.24	0.999	0.175	0.510	0.474	10.9	0.557
L61 (0.95)	2.20	1.007	0.175	0.503	0.451	11.4	0.566
L62 (0.97)	2.17	0.957	0.181	0.517	0.463	13.7	0.553
L63 (0.98)	2.16	0.894	0.191	0.521	0.448	18.1	0.528
L64 (final)	2.10	0.754	0.224	0.533	0.471	98.5	0.470

H Theoretical Predictions Across Models and Modalities

The main text validates the centroid-aligned variability theory (Sec. 5) on a handful of representative backbones. Here we apply it, unchanged, to a comprehensive list of 36 state-of-the-art models, spanning four modalities: 15 vision encoders (ImageNet), 14 autoregressive language models and 4 masked language models (TinyStories), and 3 audio encoders (VGGSound). The models have different architectures (transformers and convolutional neural networks), training methods (supervised,

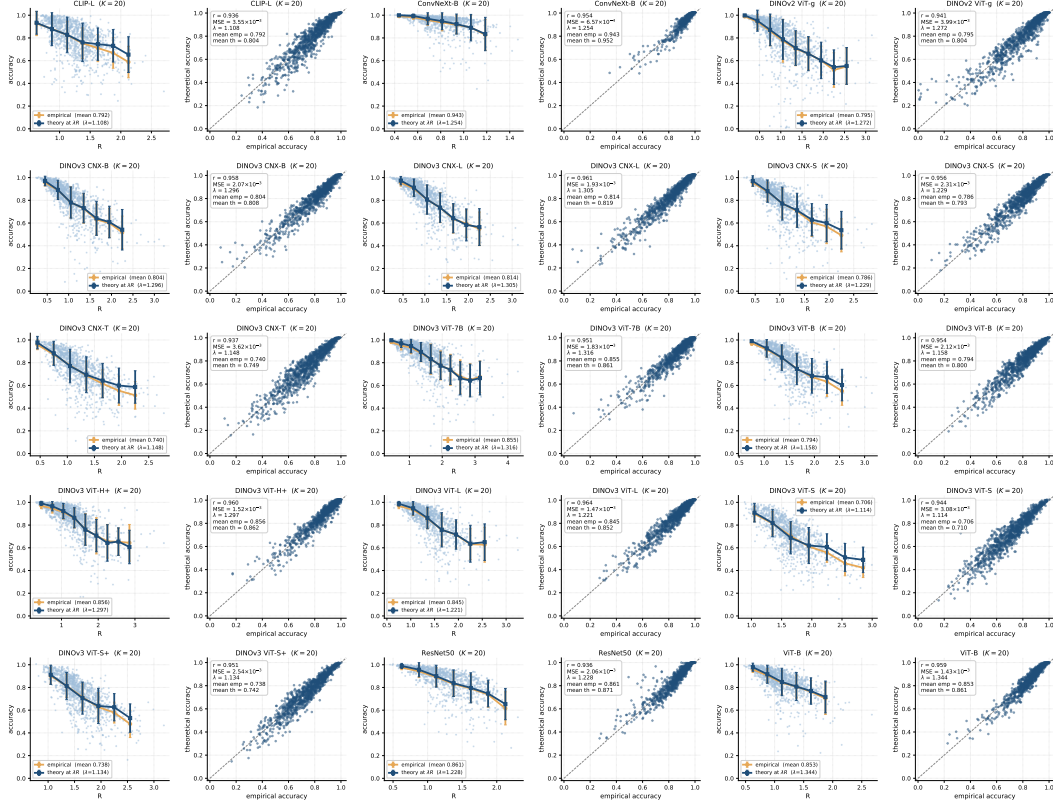


Figure 10: 15 vision models. For each backbone: (left) per-class accuracy versus empirical radius R , empirical binned mean and theory binned mean nearly overlap in most models; (right) per-class predicted versus empirical accuracy with r , RMSE and λ . The dashed line is $y = x$. Mean Pearson $r = 0.951$, mean RMSE = 0.047.

self-supervised, and contrastive), and training datasets. For every model, we compute the per-class prediction with a single fitted global rescaling λ and the measured quantities the theory predicts, and compare it to the empirical prototype accuracy class by class. No per-model tuning beyond the scalar λ is used.

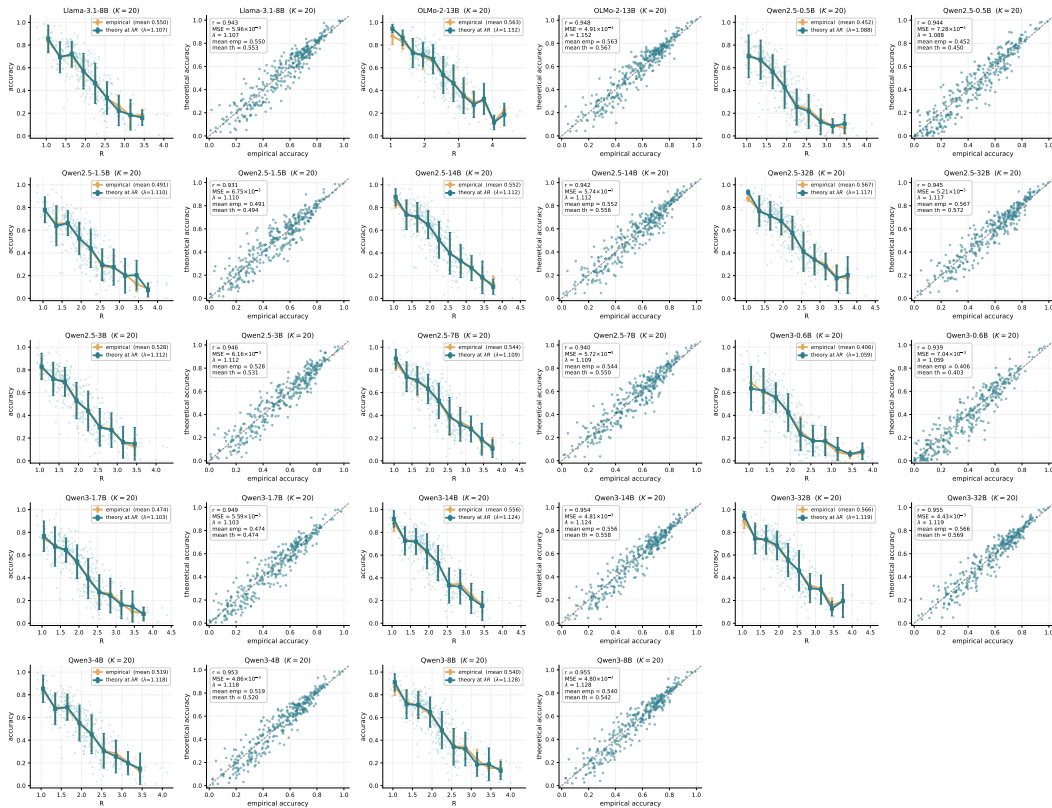
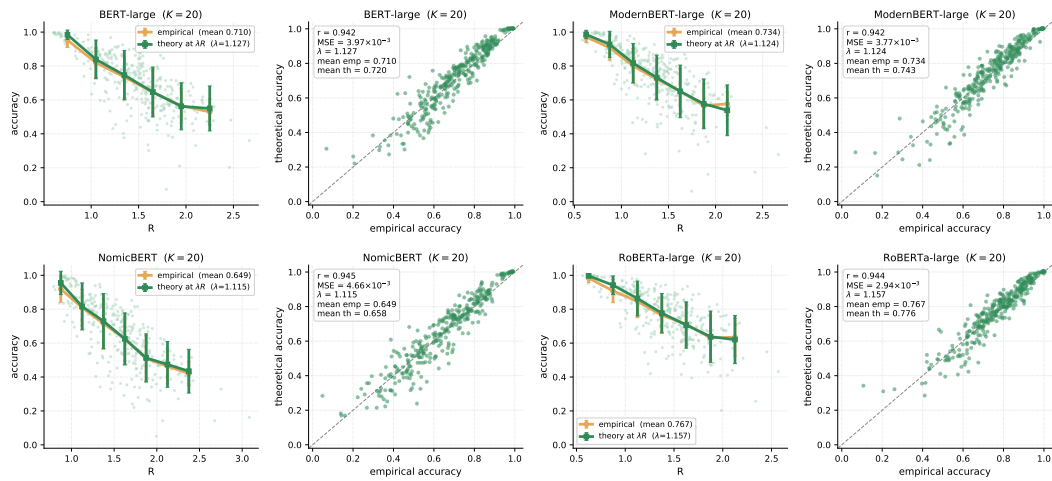
Table 12 reports, for each backbone, the number of classes P , the feature dimension N , the empirical prototype accuracy, the per-class Pearson correlation r between predicted and empirical accuracy, the fitted λ , and RMSE.

Across all four modalities the per-class correlation between the theory and the empirical accuracy is high: $r = 0.951$ (vision), 0.946 (autoregressive), 0.943 (masked) and 0.948 (audio). Per-class errors have nearly zero bias (mean per modality bias 0.002 – 0.009) and only a few percent mean deviation (mean per modality RMSE 0.047–0.076). Thus our theory predicts prototype accuracy class by class across vision, language and audio representations.

For each model we also reproduce the two per-class panels of the theory vs the empirical accuracy as in Fig. 4. The four figures show all backbones grouped by modality. The theory not only tracks the general R trend, but often closely follows the individual pattern of the model.

I Variants of the theory

The prototype-accuracy theory of the main text (Theorem 5.5) is one point in a small family of predictors, obtained by making different assumptions about the covariance of the standardized centroid projections s_μ . Here we compare the main theory against a simpler baseline and one richer refinement. All three predictors are evaluated at a fixed rival $K = 20$, on the same 36 backbones and the same per-class accuracy targets. The global rescaling factor λ was fitted separately for each variant.



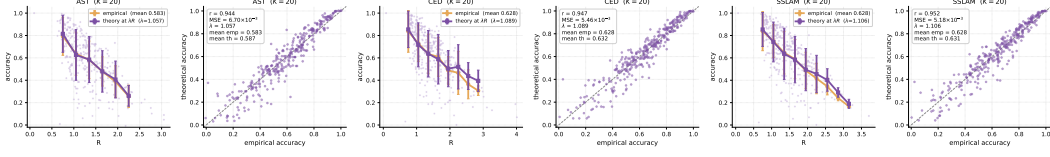


Figure 13: Audio (3 encoders, VGGSound training split, final layer). Panels as in Fig. 10. Mean $r = 0.948$, mean RMSE = 0.076.

Uncorrelated s: The model is the same as the one presented in Sec. 5, with $\rho_{\mu \neq k} = 0$, so that different s_μ are uncorrelated. The uncorrelated theory fits the data significantly worse than the theory of Sec. 5, where we consider a simple off-diagonal structure, and thus it is not used in the main text.

Cluster expansion: We start from Eq. 22. We assume that $\Sigma_{\mu\nu}$ has the structure identified in Sec. 5, plus a small correction on the terms outside the diagonal, $\Sigma = \Sigma_0 + \delta\Sigma$, $\delta\Sigma_{\mu\mu} = 0$. We shift the logits to have zero mean for convenience, and write the orthant probability conditioned on t_k , which has a diagonal element from the original structure plus the small correction

$$\text{Acc}^{\text{th}}(R) \propto \prod_{\mu \in \mathcal{R}_K} \int_{-\infty}^{t_k - f_\mu(t_k) - g_\mu} \frac{dt_\mu}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} \sum_{\mu, \nu \in \mathcal{R}_K} t_\mu (\delta_{\mu\nu} \Lambda_\mu^2 + \delta\Sigma_{\mu\nu})^{-1} t_\nu\right) \quad (29)$$

where $f_\mu(t_k) = \rho_\mu \sigma_\mu (t_k - 1) / \sigma_k$ is the conditional mean similar to Sec. C.2, and Λ_μ is the conditional standard deviation $\Lambda_\mu = R \sigma_\mu \sqrt{1 - \rho_\mu^2}$. Expanding to linear order in $\delta\Sigma_{\mu\nu}$:

$$\text{Acc}^{\text{th}}(R) = \prod_{\mu \in \mathcal{R}_K} \int_{-\infty}^{t_k - f_\mu(t_k) - g_\mu} \frac{dt_\mu}{\sqrt{2\pi} \Lambda_\mu^2} \exp\left(-\frac{1}{2} \sum_{\mu \in \mathcal{R}_K} \frac{t_\mu^2}{\Lambda_\mu^2}\right) \left(1 + \frac{1}{2} \sum_{\mu \neq \nu} \frac{\delta\Sigma_{\mu\nu} t_\mu t_\nu}{\Lambda_\mu^2 \Lambda_\nu^2}\right) \quad (30)$$

Define the conditional margin

$$F_\mu(t_k) = \frac{t_k - f_\mu(t_k) - g_\mu}{\Lambda_\mu}, \quad (31)$$

and perform the truncated Gaussian integrals

$$\text{Acc}^{\text{th}}(R) = \prod_{\mu \in \mathcal{R}_K} \Phi(F_\mu(t_k)) \left(1 + \frac{1}{2} \sum_{\mu \neq \nu} \frac{\delta\Sigma_{\mu\nu}}{\Lambda_\mu \Lambda_\nu} m(F_\mu(t_k)) m(F_\nu(t_k))\right) \quad (32)$$

We used the common inverse Mills ratio $m(x) = \phi(x)/\Phi(x)$, where $\phi(x)$ is a standard normal PDF and $\Phi(x)$ is a standard normal CDF. Substitute $t_k = Rz\sigma_k + 1$ and averaging over z we get the final result:

$$\text{Acc}^{\text{th}}(R) = \mathbb{E}_{z \sim \mathcal{N}(0,1)} \left[\prod_{\mu \in \mathcal{R}_K} \Phi(F_\mu(z)) \left(1 + \frac{1}{2} \sum_{\mu \neq \nu} \frac{\delta\Sigma_{\mu\nu}}{\Lambda_\mu \Lambda_\nu} m(F_\mu(z)) m(F_\nu(z))\right) \right] \quad (33)$$

With

$$F_\mu(z) = \frac{\sigma_k - \rho_\mu \sigma_\mu}{\Lambda_\mu / R} z + \frac{1 - g_\mu}{\Lambda_\mu} \quad (34)$$

Although the cluster expansion theory remains analytic, and fits the data slightly better than the original theory derived in Sec. 5, it adds another $K(K-1)/2$ parameters $\delta\Sigma_{\mu\nu}$ that need to be measured, and it is less transparent than the original simpler theory. Thus, we bring it here as an additional theoretical result with its performance on the 36 backbones we tested (Table. 14), but do not use it in the main text.

Backbone	Uncorrelated		Main theory		Cluster expansion	
	λ	RMSE	λ	RMSE	λ	RMSE
<i>Vision (ImageNet-1K)</i>						
ConvNeXt-B	0.979	0.0493	1.254	0.0256	1.281	0.0234
ResNet50	1.114	0.0589	1.228	0.0454	1.320	0.0434
DINOv3 ViT-H ⁺	1.212	0.0578	1.297	0.0390	1.317	0.0377
DINOv3 ViT-7B	1.248	0.0574	1.316	0.0427	1.337	0.0402
ViT-B	1.196	0.0593	1.344	0.0378	1.393	0.0339
DINOv3 ViT-L	0.996	0.0733	1.221	0.0384	1.248	0.0366
DINOv3 CNX-L	1.160	0.0706	1.305	0.0440	1.360	0.0417
DINOv3 CNX-B	1.087	0.0799	1.296	0.0455	1.337	0.0420
DINOv2 ViT-g	1.114	0.1166	1.272	0.0631	1.311	0.0561
DINOv3 ViT-B	0.864	0.0929	1.158	0.0460	1.196	0.0416
CLIP-L	0.671	0.1188	1.108	0.0596	1.199	0.0471
DINOv3 CNX-S	0.966	0.0935	1.229	0.0481	1.275	0.0450
DINOv3 CNX-T	0.855	0.1224	1.148	0.0602	1.209	0.0535
DINOv3 ViT-S ⁺	0.861	0.0981	1.134	0.0504	1.179	0.0438
DINOv3 ViT-S	0.838	0.1061	1.114	0.0555	1.173	0.0464
<i>mean</i>	1.011	0.0837	1.228	0.0468	1.276	0.0422
<i>Audio (VGGSound)</i>						
SSLAM	0.920	0.1194	1.106	0.0720	1.179	0.0555
CED	0.894	0.1164	1.089	0.0739	1.183	0.0599
AST	0.861	0.1256	1.057	0.0819	1.173	0.0617
<i>mean</i>	0.892	0.1205	1.084	0.0759	1.178	0.0590
<i>Masked LM (TinyStories, masked-word)</i>						
RoBERTa-large	0.943	0.0856	1.157	0.0542	1.235	0.0477
ModernBERT-large	0.884	0.0985	1.124	0.0614	1.225	0.0511
BERT-large	0.943	0.0930	1.127	0.0630	1.235	0.0492
NomicBERT	0.920	0.1048	1.115	0.0683	1.238	0.0565
<i>mean</i>	0.923	0.0955	1.131	0.0617	1.234	0.0511
<i>Autoregressive LM (TinyStories, next-token)</i>						
Qwen2.5-32B	0.937	0.1069	1.117	0.0722	1.363	0.0755
Qwen3-32B	0.940	0.1004	1.119	0.0665	1.376	0.0707
OLMo-2-13B	0.986	0.0922	1.152	0.0701	1.393	0.0726
Qwen3-14B	0.953	0.0984	1.124	0.0694	1.383	0.0716
Qwen2.5-14B	0.933	0.1093	1.112	0.0758	1.370	0.0773
Llama-3.1-8B	0.923	0.1061	1.107	0.0772	1.373	0.0750
Qwen2.5-7B	0.946	0.1114	1.109	0.0756	1.350	0.0780
Qwen3-8B	0.963	0.0947	1.128	0.0693	1.363	0.0703
Qwen2.5-3B	0.956	0.1058	1.112	0.0785	1.357	0.0780
Qwen3-4B	0.950	0.0993	1.118	0.0697	1.347	0.0685
Qwen2.5-1.5B	0.953	0.1166	1.110	0.0822	1.330	0.0865
Qwen3-1.7B	0.940	0.1003	1.103	0.0747	1.337	0.0718
Qwen2.5-0.5B	0.937	0.1072	1.088	0.0853	1.350	0.0808
Qwen3-0.6B	0.900	0.1096	1.059	0.0839	1.311	0.0772
<i>mean</i>	0.944	0.1042	1.111	0.0750	1.357	0.0753
Overall	0.965	0.0960	1.160	0.0618	1.295	0.0574

Table 14: Fitted global rescaling λ and per-class prediction error RMSE for three predictors at $K=20$, on all 36 backbones: the uncorrelated-projections baseline ($\rho_\mu=0$), the main theory (rank-one true-rival correlation), and the first-order cluster expansion. λ is refitted separately for each predictor.